KMT strong embedding theorems



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Coupling Binomial with Gaussian

* Let $S_n = X_1 + \ldots + X_n$ with X_k i.i.d. ± 1 with equal probability * Let Z be a standard Gaussian * CLT \Leftrightarrow Can couple so that $S_n - Z\sqrt{}$

TUSNADY'S LEMMA. For large enough n $|S_n| \leq |Z|\sqrt{n+3}$ and

* Tusnády had better constants. Irrelevant for us. Power 2 on Z important * $\mathbb{P}\{|S_n - Z\sqrt{n}| \ge x\} \le e^{-cx}$ which is far better than what CLT gave * Key ingredient in the proof of KMT theorems

$$\left| n \right| = o(\sqrt{n}) \text{ a.s.}$$

, there is a coupling so that
$$S_n - Z\sqrt{n} | \le Z^2 + 11$$

KMT for simple symmetric random walk

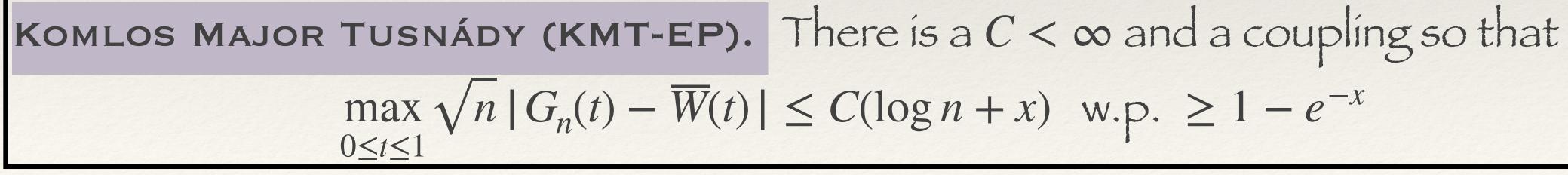
- * $S = (S_0, S_1, ...)$ SSRW as before
- * Wastandard Brownian motion in one dimension
- KOMLÓS MAJOR TUSNÁDY (KMT-RW). There is a $C < \infty$ and a coupling so that $0 \leq k \leq n$
- * Improves upon the \sqrt{n} in Donsker's theorem. Bound $\log n$ is optimal * $\log n$ as opposed to O(1) in univariate coupling * Komlós, Major, Tusnády showed this for X_i with exponential tail

* Donsker's theorem \Leftrightarrow Can couple so that $\max_{0 \le k \le n} |S_k - W_k| = o(\sqrt{n})$ a.s.

 $\max |S_k - W_k| \le C(\log n + x) \text{ w.p.} \ge 1 - e^{-x}$

KMT for uniform empirical process

- * $V_1, V_2, ... i.i.d. uniform[0,1]$
- * Empirical CDF: $F_n(t)$ = proportion of $k \le n$ for which $V_k \le t$
- * Empirical process: $G_n(t) = \sqrt{n(F_n(t) t)}$ for $0 \le t \le 1$
- * Brownian bridge: $\overline{W}(t) = W_t tW_1$ for $0 \le t \le 1$
- * Donsker (Kolmogorov-Smírnov) implies coupling so that $\max\{\sqrt{n} | G_n(t) - \overline{W}(t) | : 0 \le t \le 1\} = o(\sqrt{n}) \text{ a.s.}$



 $\max \sqrt{n} |G_n(t) - \overline{W}(t)| \le C(\log n + x) \quad \text{w.p.} \ge 1 - e^{-x}$

Broad outline of the proof

Proof strategies

Combinatorial: KMT, Csörgõ-Révész, Bretagnolle and Massart, Dudley, Massart, Carter—Pollard, Pollard (both KMT-RW and KMT-EP). Excellent reference - Pollard's book UGMPT

Analytic: Chatterjee (KMT-RW for SSRW). Extension by Bhattacharjee and Goldstein.

Steps in both kinds of proof

- involved). Analytic proof uses a form of *Stein's method*.
- that can be proved by induction on *n*.

1. Univariate coupling lemmas such as Tusnády's lemma or other similar coupling lemmas (will see in second lecture). Combinatorial proof compares Binomial to Gaussian by Stirling's etc. (but way more

2. Go from univariate coupling to coupling of paths. In the dyadic approach, couple ($G_n(1/2), \overline{W}(1/2)$), $(G_n(1/4), \overline{W}(1/4)), (G_n(3/4), \overline{W}(3/4)), \dots$ for log *n* generations. Chatterjee craftily frames a statement



An idea: Use the Cauchy criterion

* To show $x_n \to x$, enough to show $x_n \to x$ $|x_n - x| \le \sum_{j \ge n} |x_j - x_{j+1}| \lesssim \frac{1}{n} \text{ or } |x_j|$

show closeness of $2S_n$ to S_{4n} (like with like).

* As $S_m \sim$ Binomial, hope for purely combinatorial proof (this lecture) or using finite Markov chains (second lecture)

$$\begin{aligned} x_n - x_{n+1} &| \le \frac{1}{n^2} \text{ or } |x_n - x_{2n}| \le \frac{1}{n^{\varepsilon}}. \text{ Then} \\ x_n - x &| \le \sum_{j \ge 0} |x_{n2^j} - x_{n2^{j+1}}| \lesssim \frac{1}{n^{\varepsilon}}. \end{aligned}$$

* Analogously, to show closeness of S_n to $Z\sqrt{n}$ (like with unlike), suffices to



Our version of the combinatorial proof

Tusnády-like lemma

TUSNADY TYPE LEMMA. For large

 $|2S_n| \le |S_{4n}| + 2$ an

Deducing Tusnády's lemma

Fix a large enough even number n and let

Successively couple so that $|Z_k| \leq |Z_{k+1}|$

* By Bernstein/Hoeffding, $|Z_k| \le 2^{k/10}$ with high probability. * Hence, a.s., $\{Z_k\}$ is Cauchy and $Z_k \rightarrow Z \sim N(0,1)$ Summing, $|Z_k| \le |Z| + \frac{2}{\sqrt{n}}$ and $|Z_k - Z| \le \frac{1}{4\sqrt{n}}Z^2 + \frac{9}{\sqrt{n}}$ which is Tusnády's lemma (for even *n*)

e enough even *n*, there is a coupling so that

$$|2S_n - S_{4n}| \le \frac{|S_{4n}|^2}{8n} + 9$$

$$Z_{0} = \frac{S_{n}}{\sqrt{n}}, Z_{1} = \frac{S_{4n}}{\sqrt{4n}}, \dots, Z_{k} = \frac{S_{4^{k}n}}{\sqrt{4^{k}n}}, \dots$$
$$+ \frac{2}{2^{k}\sqrt{n}} \text{ and } |Z_{k} - Z_{k+1}| \le \frac{1}{2^{k+2}\sqrt{n}} |Z_{k+1}|^{2} + \frac{9}{2^{k+1}\sqrt{n}}$$

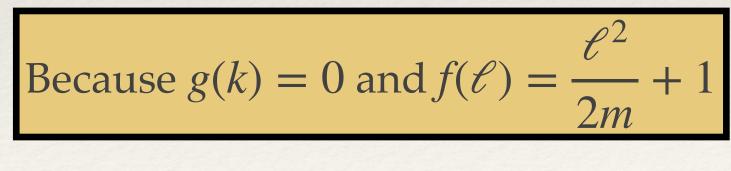


How to couple $2S_n$ and S_{4n} ? Preparation

FACT (easy to prove)

 $X, Y \sim \mathbb{Z}_+ \text{-valued. Assume } \mathbb{P}\{X \ge k - f(k)\} \ge \mathbb{P}\{Y \ge k\} \text{ and } \mathbb{P}\{Y \ge k - g(k)\} \ge \mathbb{P}\{X \ge k\} \text{ for all } k.$ Then, they can be coupled so that $X \ge Y - f(Y)$ and $Y \ge X - g(X)$.

* Towards the goal of coupling
$$2S_n$$
 with S_{4n} , let $n = 2m$.
Let $\alpha_m(k) = \mathbb{P}\{S_{2m} = 2k\}$ and $\beta_m(k) = \mathbb{P}\{S_{8m} = 4k \text{ or } 4k - 2\}$. PMFs of $X = \frac{S_{2m}}{2}$ and $Y = \left[\frac{S_{8m}}{4}\right]_+$
Let $\overline{\alpha}_m(k) = \sum_{j \ge k} \alpha(j)$ and $\overline{\beta}_m(k) = \sum_{j \ge k} \beta_m(j)$ be the tails
* From the fact, it suffices to prove that
 $\overline{\alpha}_m(k) \le \overline{\beta}_m(k)$ for all k and $\overline{\alpha}_m(k) \ge \overline{\beta}_m(\ell)$ for $k \le \ell - \frac{\ell^2}{2m} - 1$



$$\ell) \text{ for } k \leq \ell - \frac{\ell^2}{2m} - \frac{\ell^2}{2m}$$



How to couple $2S_n$ and S_{4n} ? Reduction

$$\overline{\alpha}_{m}(k) = \sum_{j \ge k} \alpha(j) \text{ where } \alpha_{m}(k) = \mathbb{P}\{S_{2m} = 2k\} = \binom{2m}{m+k} \frac{1}{2^{2m}}$$

$$\overline{\beta}_{m}(k) = \sum_{j \ge k} \beta_{m}(j) \text{ where } \beta_{m}(k) = \mathbb{P}\{S_{8m} = 4k \text{ or } 4k-2\} = \binom{8m+1}{4m+2k} \frac{1}{2^{8m}}$$

$$\overline{\beta}_{m}(k) = \sum_{j \ge k} \beta_{m}(j) \text{ where } \beta_{m}(k) = \mathbb{P}\{S_{8m} = 4k \text{ or } 4k-2\} = \binom{8m+1}{4m+2k} \frac{1}{2^{8m}}$$

$$\overline{\beta}_{m}(k) = \sum_{j \ge k} \beta_{m}(j) \text{ where } \beta_{m}(k) \le \overline{\beta}_{m}(k) \text{ for all } k \text{ and } \overline{\alpha}_{m}(k) \ge \overline{\beta}_{m}(\ell) \text{ for } k \le \ell - \frac{\ell^{2}}{2m} - 1$$

$$\overline{\beta}_{m}(k) = \sum_{j \ge k} \alpha_{m}(\ell) \text{ for } k \le \ell - \frac{\ell^{2}}{2m} - 1$$

$$\overline{\beta}_{m}(k) = \frac{\sqrt{n}}{\sqrt{k(n-k)}} e^{-nD(k/n)} \le \frac{1}{2^{n}} \binom{n}{k} \le B \frac{\sqrt{n}}{\sqrt{k(n-k)}} e^{-nD(k/n)}$$

$$\overline{\beta}_{m}(k) = \frac{\sqrt{n}}{2m} \frac{1}{2^{n}} \binom{n}{k} \le e^{-nD(k/n)}$$

$$\overline{\beta}_{m}(k) = \frac{\sqrt{n}}{2m} \frac{1}{2^{n}} \frac{1}{2^{n}} \binom{n}{k} \le e^{-nD(k/n)}$$

pg(2-2p)

he far tail





How to couple $2S_n$ and S_{4n} ? The key lemma

KEY COMBINATORIAL LEMMA. $\alpha_m(k) \le \beta_m(k)$ for all k and $\alpha_m(k) \ge \beta_m(\ell)$ for $k \le \ell - \frac{\ell^3}{4m^2} - 1$

* Compares mass functions instead of tails * Must be combined with estimates for binomial coefficients to complete the proof

* Bíjective proof? (see next slide)

* As $\frac{\ell^3}{m^2} \lesssim \frac{\ell^2}{m}$ it suggests Carter-Pollard improvement: $\frac{1}{n} |Z|^3 \leq Z^2$ in Tusnády's lemma bound





Proof of the key combinatorial lemma

KEY COMBINATORIAL LEMMA. $\alpha_m(k) \leq \beta_m(k)$

Lemma follows from following claims

 $f(m, k) \ge 1$ 1. $f(m, k + 1) \ge f(m, k)$ for $1 \le k \le m - 1$ and $m \ge 1$ 2. $f(m, 1) \ge f(m + 1, 1)$ for $m \ge 1$ 3. $f(m, 1) \to 1$ as $m \to \infty$

k) for all k and
$$\alpha_m(k) \ge \beta_m(\ell)$$
 for $k \le \ell - \frac{\ell}{4m}$

Let
$$f(m,k) := \frac{\beta_m(k)}{\alpha_m(k)} = \frac{\binom{8m+1}{4m+2k}}{2^{6m}\binom{2m}{m+k}}$$
 and $g_h(m,k) := \frac{\beta_m(k)}{\alpha_m(k-h)} = \frac{\binom{8m+1}{4m+2k}}{2^{6m}\binom{2m}{m+k-h}}$

$$g_{h}(m,k) \leq 1 \text{ if } h + 1 \leq k \leq [(4h - 1)m^{2}]$$
1. $g_{h}(m,k+1) \leq g_{h}(m,k) \text{ for } h + 1 \leq k \leq [(4h - 1)m^{2}]$
2. $g_{h}(m,h+1) \geq g_{h}(m+1,h+1) \text{ for } m \geq h+1$
3. $g_{h}(m,h+1) \rightarrow 1 \text{ as } m \rightarrow \infty \text{ for } h \text{ fixed}$





Proof of the key combinatorial lemma

Proof that $f(m, k) \ge 1$

 $f(m,1) \rightarrow 1$ as $m \rightarrow \infty$ is trivial (Stirling's or invoke local limit theorem).

In the other two steps, we get to take ratios and cancel most of the factorials

To show that $f(m,1) \ge f(m + 1,1)$, compute f(m + 1,1)/f(m,1) to get

 $m(m+2) \times (8m+2)...(8m+9)$

 $2^{6}(2m+1)(2m+2) \times (4m+3)...(4m+6) \times (4m+6)$

where x = 8m and p(x) = (x + 16)(x + 9)(x + 7)(x + 5)(x + 3)q(x) = (x + 12)(x + 10)(x + 8)(x + 6)(x + 4)

(16,9,7,5,3) majorizes (12,10,8,6,4), hence $p(x) \le q(x)$ for x > 0 by Schur concavity

$$\frac{1}{4m}\dots(4m+3) = \frac{p(x)}{q(x)}$$

$$f(m,1) = \frac{\binom{8m+1}{4m+2}}{2^{6m} \binom{2m}{m+1}}$$

Majorization explained 16>12 16+9=25>22=12+10 16+9+7=**32**>**30**=12+10+8 16+9+7+5=37>36=12+10+8+6 16+9+7+5+3=40=12+10+8+6+4



Proof of the key combinatorial lemma

Remains to show that $f(m, k + 1) \ge f(m, k)$ for $1 \le k \le m - 1$. Again take ratios $\frac{f(m,k+1)}{f(m,k)} = \frac{(x+2a+4)(x-a)(x-a+1)}{(x-2a)(x+a+1)(x+a+2)} \text{ with } x = 4m, a = 2k$ Do not see any majorization, but Numerator-Denominator gives $2x(x - 2a) + 2x + 4a^{2}(a + 2) \ge 0$ since x - 2a = 4(m - k) > 0

Remark. It is in this step for the second claim that we get $g_h(m, k + 1) \leq g_h(m, k)$ provided $2((4h-1)x^2 - 2a^3) + 2x(8h+2a-1) + 8(a^2(h-1) + ah + h) \ge 0$ Second and third terms are positive. For the first to be positive, require $k^3 \leq (4h - 1)m^2$



Our version of the analytic proof

KMT for simple symmetric random walk

- * $S = (S_0, S_1, ...) SSRW$
- * Wastandard Brownian motion in one dimension
- KOMLÓS MAJOR TUSNÁDY (KMT-RW). There is a $C < \infty$ and a coupling so that $0 \leq k \leq n$
- * Improves from \sqrt{n} in Donsker's theorem to log *n* which is optimal * Komlós, Major, Tusnády showed this for X_i with exponential tail

* Donsker's theorem \Leftrightarrow Can couple so that $\max_{0 \le k \le n} |S_k - W_k| = o(\sqrt{n})$ a.s.

 $\max |S_k - W_k| \le C(\log n + x) \text{ w.p.} \ge 1 - e^{-x}$

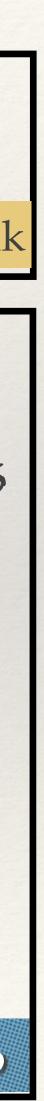
Chatterjee's approach to KMT for SSRW

lemmas coupling binomial and hypergeometric to Gaussian.

From univariate coupling to that of paths Induction on n (length of the random walk) applied to the following statement: For any probable value t of S_n , there is a coupling of the random walk conditioned to have $S_n = t$ (a random walk bridge) with Brownian motion conditioned to have W(n) = t (Brownian bridge) such that the maximum difference $M_{n,t}$ between the two paths satisfies $\mathbb{E}[e^{\lambda M_{n,t}}] \leq e^{A\log n + B\lambda^2 \frac{t^2}{n}} \text{ for } \lambda \leq \lambda_0.$ Here λ_0, A, B are fixed constants.

Univariate coupling lemmas In place of Tusnády's lemma, Chatterjee proves two Subject of the rest of the talk

We do not talk about this second step



Coupling lemma for Binomial

with equal number of coupons labelled + 1 and - 1

* Like in Tusnády's lemma, two random variables of std. dev. \sqrt{n} are coupled within unit distance * Tusnády's lemma implies this because $\mathbb{E}[e^{\theta_0 Z^2}] < \infty$ for small enough θ_0 * Converse not possible, as largeness of the difference not related with the value of Z (or S_n)

- * Let $S_k = X_1 + ... + X_k$ where X_j are i.i.d. ± 1 . SRSWR of k coupons from a box
- CHATTERJEE'S BINOMIAL COUPLING LEMMA. For some θ_0 , κ_0 , there is a coupling for any $n \geq 1$ such that $\mathbb{E}\left[e^{\theta_0|S_n-Z\sqrt{n}|}\right] \leq \kappa_0.$



Coupling lemma for hypergeometric

* Let $S_k[n, s]$ be the sum of k coupons drawn without replacement from a box containing n coupons labelled ± 1 whose sum is s i.e., $(n \pm s)/2$ coupons of either kind. (Hypergeometric)

* Let
$$W_k[n,s] = S_k[n,s] - \frac{k}{n}s$$
 be the centred ve

* Let
$$\sigma_{n,k}^2 = \frac{k(n-k)}{n-1}$$
 be the variance of $W_k[n, 0]$

for
$$n \ge 1$$
 and $\frac{1}{3}n \le k \le \frac{2}{3}n$ such that $\mathbb{E}\left[e^{n}\right]$

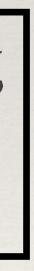
* Supplements binomial coupling and feeds into the induction step (which is about bridges)

* Enough to have
$$k \approx \frac{n}{2}$$
 (nearest integer, for

- version. Its variance is $\left(1 \frac{s^2}{n^2}\right) \frac{k(n-k)}{n-1}$
- 0] (case of unbiased box)

CHATTERJEE'S HYPERGEOMETRIC COUPLING LEMMA. For some θ_1, M_1 , there is a coupling $\theta|W_k[n,s] - \sigma_{n,k}Z| \le e^{1 + M_1 \theta^2 \frac{s^2}{n}} \text{ for } \theta \le \theta_1.$





Univariate couplings using Cauchy criterion

From the following lemmas, it is easy to deduce Chatterjee's coupling lemmas, using the Cauchy criterion as in the first part of the talk

BINOMIAL COUPLING LEMMA. For some θ_0, κ_0 , there is a coupling for any $n \ge 1$ such that $\mathbb{E}\left[e^{\theta_0|2S_n-S_{4n}|}\right] \leq \kappa_0$

with
$$W_1 = W_k[n,0], W_2 = W_{4k}[4n,0], W = W_k$$

(1) $\mathbb{E}\left[e^{\theta_1|2W_1 - W_2|}\right] \le \kappa_1$
(2) $\mathbb{E}\left[e^{\theta|W_1 - W|}\right] \le e^{1 + M_1 \theta^2 \frac{s^2}{n}}$ for any $\theta \le \theta_1$

HYPERGEOMETRIC COUPLING LEMMA. For some θ_1, κ_1, M_1 , for even n and $\frac{1}{3}n \le k \le \frac{2}{3}n$, $Y_k[n,s] - \frac{sk}{m}$, there are couplings such that



Coupling distributions on Z using Markov chains



The general strategy of coupling

- * α , β probability distributions on \mathbb{Z} that we want to couple. Eg., $2S_n$, S_{4n}
- * We shall need nearest neighbour Markov chains on \mathbb{Z} with these stationary distributions. This means no gaps in support
- ★ Examples: $\frac{1}{2}S_{4n}$ has support $[-2n,2n] \cap \mathbb{Z}$ Good! But S_n has gaps. Modify to $S_n + R$ where R is an independent r.v. taking values -1,0,1 w.p. $\frac{1}{4}, \frac{1}{2}, \frac{1}{4}$. Enough to couple $S_n + R$ with $\frac{1}{2}S_{4n}$
- * Similar games (scaling, additive perturbation) for hypergeometric coupling



The general strategy of coupling

- Assume that α has finite connected support in \mathbb{Z} . Many choices of rates λ_i^{\pm} for $i \mapsto i \neq 1$ *
- **Ehrenfest-like chain**: If α is reversible for $\lambda_i^{\pm} = T(i) \mp i$ for some $T : \mathbb{Z} \mapsto \mathbb{R}$ *
- Does it exist? The equations $\alpha(i)(T(i) i) = \alpha(i + 1)(T(i + 1) + i + 1)$ can be solved from right end of * support to get $T(i) = i + \frac{2}{\alpha(i)} \sum_{i > i} j\alpha(j)$
- Satisfied at left end if and only if α has **zero mean** *

T is called the Stein coefficient of α

Example: If α is the distribution of $S_{4n}/2$, then T(i) = 2n for $-2n \le i \le 2n$ (The true Ehrenfest chain) **Example:** If β is the distribution of $S_n + R$, then $T(j) = \begin{cases} 2n+1 & \text{if } j \in \{-n, -n+2, \dots, n-2, n\} \\ 2n+2 - \frac{j^2}{n+1} & \text{if } j \in \{-n-1, -n+1, \dots, n-1, n+1\} \end{cases}$

if
$$j \in \{-n, -n+2, ..., n-2, n\}$$

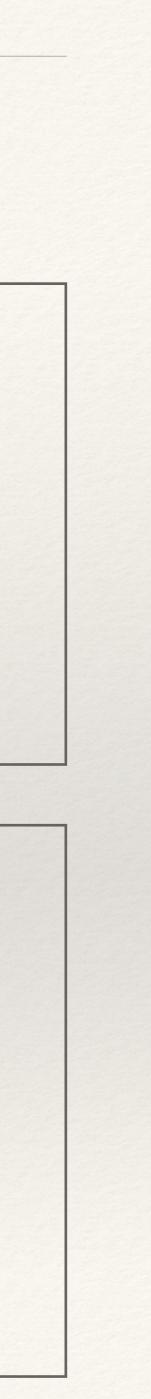
Relationship to Stein's method

- * Stein's method: $\mathbb{E}[Wf(W)] = \mathbb{E}[\sigma^2 f'(W)]$ for a large class of *f* if and only if $W \sim N(0, \sigma^2)$
- distribution $\pi = N(0,1)$. Stein's equation is $\mathbb{E}_{\pi}[Lg] = 0$, with f = g'
- coefficient of W. Closeness of T to constant related to closeness of W to normal.
- * and only if $W \sim S_n$
- distribution $\pi = \mathscr{L}(S_n)$. Stein's equation is $\mathbb{E}_{\pi}[Lh] = 0$, with h(i) = f(i) f(i-2)
- coefficient of W. Closeness of T to constant related to closeness of W to normal.
- Equivalently, $\mathscr{L}(W)$ is stationary for a Markov chain with rates $\lambda_i^{\pm} = T(i) \mp i$ *

* The Ornstein-Uhlenbeck process has generator Lg(x) = g''(x) - xg'(x) and stationary * If a random variable W satisfied $\mathbb{E}[Wf(W)] = \mathbb{E}[Tf'(W)]$, Chatterjee calls T the Stein

Stein's equation for Binomial: $\mathbb{E}[n(g(W+2) - g(W))] = \mathbb{E}[W(g(W+2) - g(W))]$ for a large class of g if

* Ehrenfest chain has generator Lg(i) = (n - i)(g(i + 2) - g(i)) + (n + i)(g(i - 2) - g(i)) and stationary * If a random variable W satisfied $\mathbb{E}[T(g(W+2) - g(W))] = \mathbb{E}[W(g(W+2) - g(W))]$, we call T the Stein



Coupling nearest neighbour Markov chains on Z

- * Let α , β be stationary distributions for rates λ_i^{\pm} , μ_i^{\pm} (n.n, continuous times chains on \mathbb{Z})
- * possible. Never move in opposing directions
- Generator $Lf(i,j) = \theta_{i,j}^{+,+}(f(i+1,j+1) - f(i,j)) + \theta_{i,j}^{+,\circ}(f(i+1,j+1) - f(i,j+1) - \theta_{i,j}^{+,\circ}(f(i+1,j+1) - \theta_{i,j}) + \theta_{i,j}) + \theta_{i,j}^{+,\circ}(f(i+1,j+1) - \theta_{i,j}) + \theta_{i,j}) + \theta_{i,j}^{+,\circ}(f(i+1,j+1) - \theta_{i,j}) + \theta_{i,j}) + \theta_{i,j}) + \theta_{i,j}) + \theta_{i,j}^{+,\circ}(f(i+1,j+1) - \theta_{i,j}) + \theta_{i,$ $+\theta_{i,i}^{-,-}(f(i-1,j-1)-f(i,j))+\theta_{i,i}^{-,\circ}(j)$

$$\theta_{i,j}^{-,\circ} = (\lambda_i^- - \mu_j^-)_+$$

$$\theta_{i,j}^{-,-} = \lambda_i^- \wedge \mu_j^- \qquad \theta_{i,j}^{\circ,-}$$

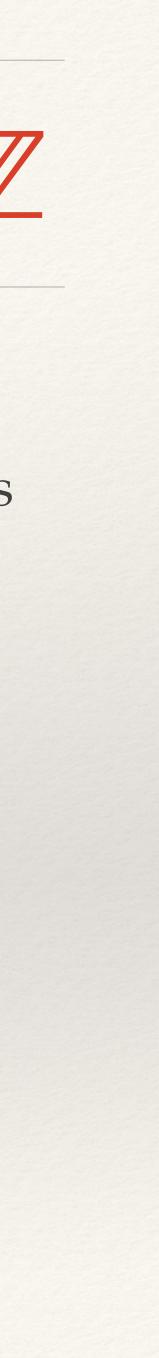
Define a Markov chain on \mathbb{Z}^2 with rates as follows. Basic idea: Try to move together as much as

$$f(i+1,j) - f(i,j)) + \theta_{i,j}^{\circ,+}(f(i,j+1) - f(i,j))$$

$$f(i-1,j) - f(i,j)) + \theta_{i,j}^{\circ,-}(f(i,j-1) - f(i,j))$$

 $\theta_{i,j}^{\circ,+} = (\mu_j^+ - \lambda_i^+)_+ \quad \theta_{i,j}^{+,+} = \lambda_i^+ \wedge \mu_j^+$ $(i,j) \qquad \qquad \theta_{i,j}^{+,\circ} = (\lambda_i^+ - \mu_j^+)_+$

 $\mu_i^- = (\mu_i^- - \lambda_i^-)_+$



Coupling nearest neighbour Markov chains on Z

- * Let Z = (X, Y) denote the Markov chain on \mathbb{Z}^2 $Lf(i,j) = \theta_{i,j}^{+,+}(f(i+1,j+1) - f(i,j)) + \theta_{i,j}^{+,\circ}(f(i+1,j+1) - f(i,j+1) - \theta_{i,j}) + \theta_{i,j}^{+,\circ}(f(i+1,j+1) - \theta_{i,j}) + \theta_{i,j}) + \theta_{i,j}^{+,\circ}(f(i+1,j+1) - \theta_{i,j}) + \theta_{i,j}^{+,\circ}(f(i+1,j+1) - \theta_{i,j}) + \theta_{i,j}) + \theta_{i,j}^{+,\circ}(f(i+1,j+1) - \theta_{i,j}) + \theta_{i,j}) + \theta_{i,j}) + \theta_{i,j}^{+,\circ}(f(i+1,j+1) - \theta_{i,j}) + \theta_{i$ $+\theta_{i,i}^{-,-}(f(i-1,j-1)-f(i,j))+\theta_{i,i}^{-,\circ}(f(i-1,j-1))$
- * If γ is a stationary distribution of Z, then $\mathbb{E}_{\gamma}[Z]$
- * $f(x,y) = \varphi(x)$: $Lf(i,j) = \lambda_i^+(\varphi(i+1) \varphi(i))$
- * $f(x, y) = \psi(y)$: $Lf(i, j) = \mu_i^+(\psi(j+1) \psi(j))$
- * Let H = X Y. If $f(x, y) = \varphi(x y)$ then $\mathbb{E}_{\gamma}[A$
 - $\mathbb{E}_{\gamma}[(B |A|)(\psi(H) \psi(H 1)]]$ $\psi(h) = \varphi(h + 1) \varphi(h) \qquad A(i, j) = \frac{1}{2}(\lambda_i^+ \mu)$

$$Z^{2} \text{ with generator}$$

$$(i + 1,j) - f(i,j)) + \theta_{i,j}^{\circ,+}(f(i,j+1) - f(i,j))$$

$$f(i - 1,j) - f(i,j)) + \theta_{i,j}^{\circ,-}(f(i,j-1) - f(i,j))$$

$$(f(Z)] = 0 \text{ for all } f : \mathbb{Z}^{2} \mapsto \mathbb{R}$$

$$+ \lambda_{i}^{-}(\varphi(i-1) - \varphi(i)) \Longrightarrow \text{First marginal of } \gamma \text{ is } \alpha$$

$$+ \mu_{j}^{-}(\psi(j-1) - \psi(j)) \Longrightarrow \text{Second marginal of } \gamma \text{ is } \mu$$

$$Lf(Z)] = 0 \text{ reduces to}$$

$$0] = 2\mathbb{E}_{\gamma}[A_{-}\psi(H-1) - A_{+}\psi(H)] - (\#)$$

$$\mu_{j}^{+} - \lambda_{i}^{-} + \mu_{j}^{-}) \quad B(i,j) = \frac{1}{2}(|\lambda_{i}^{+} - \mu_{j}^{+}| + |\lambda_{i}^{-} - \mu_{j}^{-}|)$$



Coupling Ehrenfest-like chains

- * Ehrenfest-like chains: $\lambda_i^{\pm} = S(i) \mp i$ and μ_i^{\pm} A(i, j) = j - i and B(i, j) = [i]
- Writing Q = |S(X) T(Y)|, (#) becomes

* Use? Put $\psi(x) = \mathbf{1}_{x \ge a}$ to get $\mathbb{E}_{\gamma}[H_+ \mathbf{1}_{H \ge a+1}] \le \mathbb{E}_{\gamma}[(Q - a)_+]$ or $\mathbb{P}_{\gamma}[H \ge a + 1] \le \mathbb{E}_{\gamma}[(Q - a)_+]$

Example: $X = \frac{1}{2}S_{4n}$ and $Y = S_n + R$ where $R \sim \frac{1}{2}S_2$ is an independent copy. Then $S(i) = 2n \text{ for } |i| \le n \quad \text{and } T(j) = \begin{cases} 2n+1 & \text{if } j \in \{-n, -n+1, \dots, n-1, n\} \\ 2n+2 - \frac{j^2}{n+1} & \text{if } i \in \{-n-1, -n+1, \dots, n-1, n+1\} \end{cases}$ $\frac{1}{n+1}$ which has exponential tail. Immediately gives *Binomial coupling lemma*. Here $Q \le 2 + -$

$$= T(j) \mp j. \text{ Then}$$

$$i - j| + (|T(j) - S(i)| - |j - i|)_{+}$$

 $\mathbb{E}_{\gamma}[(Q - |H|)_{+}(\psi(H) - \psi(H - 1))] = 2\mathbb{E}_{\gamma}[H_{+}\psi(H - 1) - H_{-}\psi(H)]$



Coupling Ehrenfest-like chains

- General bound. With Q = |S(X) T(Y)| and H = X Y, for any function ψ * $\mathbb{E}_{\gamma}[(Q - |H|)_{+}(\psi(H) - \psi(H - 1))] = 2\mathbb{E}_{\gamma}[H_{+}\psi(H - 1) - H_{-}\psi(H)]$
- constants *K*, κ , Θ (the quadratic θ^2 in the exponent is important)
- * **Proof of (*):** Taking $\psi(x) = e^{\theta(x+1)}\mathbf{1}_{x>0}$ and $\psi(x) = e^{\theta(-x+1)}\mathbf{1}_{x<0}$ and add to get $2\mathbb{E}_{\gamma}[|H|e^{\theta|H|}] \le (e^{\theta} - 1)\mathbb{E}_{\gamma}[Qe^{\theta|H|}]$ $\leq (e^{\theta} - 1) \left\{ \beta \theta \mathbb{E} \left[|H| e^{\theta |H|} \right] + \beta e^{-1} \mathbb{E} \left[e^{Q/\beta} \right] \right\}$

because $\beta x \log x$ and $\beta e^{-1+\frac{x}{\beta}}$ are Legendre convex duals of each other. Choose $\beta = 1/\theta^2$ and rearrange. Use for small enough θ to get $\mathbb{E}_{\gamma}[\theta | H | e^{\theta | H |}] \leq K \mathbb{E}_{\gamma}[e^{\kappa \theta^2 Q}]$. Now use $e^x \leq e^b + \frac{1}{2}xe^x$ with any b. D

* Stein coefficients)

Exponential tail on H not good enough for hypergeometric coupling lemma. Better bound: For some universal

 $\mathbb{E}\left[e^{\theta|H|}\right] \leq K \mathbb{E}\left[e^{\kappa\theta^2 Q}\right] \text{ for } \theta \leq \Theta \quad -(*)$

If *f* is convex, its dual is $f^{\dagger}(y) := \sup\{xy - f(x)\}$ Clearly $\forall x, y$, $xy \le f(x) + f^{\dagger}(y)$

The essential point is this: In the Markov coupling, H = X - Y is controlled by Q = |S - T| (the difference in



Binomial coupling

* Let $S_{n,p} \sim \text{Binomial}(n,p)$ (centered) taking values in $\{-np, -np + 1, ..., nq - 1, nq\}$ * It has Stein coefficient T(x) = 2pqn + (q - p)x* $X = S_{4n,\frac{1}{2}}$ has $T_X(x) = 2n$ and $Y = 2S_{n,\frac{1}{2}} + R$ has $T_Y(y) = 2n + O(y^2/n)$. Thus $Q \leq -$ in this case. Gives the Binomial coupling lemma. * $X = S_{n,\frac{1}{2}}$ has $T_X(x) = n/2$ and $Y = S_{n,p}$ has T $Q \leq \frac{n}{2} |1 - 4pq| + |p - q| Y$ which is bour

* In the Markov chain coupling $\mathbb{E}[e^{\theta |H|}] \leq \mathbb{E}[e^{\kappa \theta^2 Q}] \leq e^{c\theta^2 \frac{s^2}{n}}$

$$T_Y(y) = 2pqn + (q - p)y$$
. Thus
indeed by $\frac{s^2}{n} + \frac{s}{\sqrt{n}} \frac{|Y|}{\sqrt{n}}$ with $s = n(p - q)$

Hypergeometric coupling lemma

HYPERGEOMETRIC COUPLING LEMMA. For with $W_1 = W_k[n,0], W_2 = W_{4k}[4n,0], W = W_k$ (1) $\mathbb{E}\left[e^{\theta_1|2W_1-W_2|}\right] \leq \kappa_1$ (2) $\mathbb{E}\left[e^{\theta|W_1-W|}\right] \le e^{1+M_1\theta^2\frac{s^2}{n}}$ for any $\theta \le \theta_1$

* Can compute explicitly the Stein coefficients of W_1, W_2, W $\mathbb{E}[f(W)] \leq \mathbb{E}[f(S_{n,p})]$ for any convex function f

some
$$\theta_1, \kappa_1, M_1$$
, for even n and $\frac{1}{3}n \le k \le \frac{2}{3}n$,
 $K_k[n,s] - \frac{sk}{n}$, there are couplings such that

* Entirely analogous to the previous situation $(W_1, W_2, W) \Leftrightarrow (S_{n,\frac{1}{2}}, S_{4n,\frac{1}{2}}, S_{n,p})$ with s = n(p-q)

* Bounding exponential moments of Q made easy by a lemma of Hoeffding that says





Thank you for listening!