



# Active Infections and Disease Extinction in the Stochastic SIR Model

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# Cheat Sheet

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# One Slide Summary

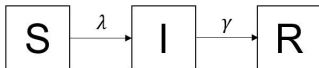
- SIR models are useful for studying the **spread of infections**.
- Our work concerns a **discrete-time SIR model** and focuses on:
  1. *Active Infections*: Confirmed cases minus recoveries.
  2. *Disease Extinction Time*: Active infections become zero.
- Our analysis via novel **stopping times** is fundamentally different.
- Our results include:
  1. Bounds for *active infections* and *disease extinction time*.
  2. Estimate for the expected value of the *largest epidemic size*.

# Background

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# SIR Models: Preliminaries

- The SIR model is an example of a [compartment](#) model.
- A fixed size population is divided into different compartments.



- Arrows indicate how the status of an individual may change.
- $\lambda$  and  $\gamma$  are parameters that control the rate of evolution.
- Two categories of SIR models: [deterministic](#) and [stochastic](#).

# Deterministic SIR Models

- First proposed by Kermack and McKendrick in 1927.
- Good if community is **homogeneous** and people **mix uniformly**.
- Described by **ordinary differential equations**:

$$\dot{S}(t) = -\lambda \frac{SI}{n}, \quad \dot{I}(t) = \lambda \frac{SI}{n} - \gamma I, \quad \dot{R}(t) = \gamma I,$$

where  $n$  is the fixed population size.

- Two major issues:
  1. **No analytical estimates** exist for **active infections**, i.e.,  $I(t)$ .
  2. **Early disease termination** not possible as  $I(t) \neq 0$  for finite  $t$ .

# When are deterministic models insufficient?

- Inherent uncertainty in epidemic
  - Community is small, e.g., school.
- Epidemic fails to start
  - Large community, but outbreak started by few individuals.
- Noisy data
  - Disease outbreak data has standard measurement errors.
- Early disease extinction
  - Such questions are of interest.

# Stochastic SIR Models

- Described via [discrete/continuous time Markov chains](#) or stochastic differential equations.
- Major issue with Markov chain based models:  
[Complicated analysis](#) via [multiple approximations](#),  
e.g., branching process for early and final stages, ODE for middle.



## Our Model

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# Continuous-time Stochastic SIR Model

- Introduced by Barlett in 1949.
- Fixed population size:  $n$
- There is an independent  $\exp(\lambda/(n-1))$  clock for each pair. Each time a clock ticks, the corresponding pair meet.
- Mean waiting before individual  $i$  meets another person is  $1/\lambda$ .
- An infected person recovers in  $\exp(\gamma)$  time, independently.
- Basic reproduction number  $\mathcal{R}_0 = \frac{\lambda}{\gamma} = \frac{\text{mean recovery time}}{\text{mean waiting time}}$ .

# Discrete-time Stochastic SIR Model

- At jump  $t \geq 0$ ,

$S_t$ : Susceptibles	$I_t$ : Infected	$R_t$ : Recovered
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- $I_t = 0$  implies  $I_{t+1} = 0$ ,  $S_{t+1} = S_t$ , and  $R_{t+1} = R_t$ .
- Suppose  $I_t > 0$ . Then, conditioned on the value of  $(S_t, I_t, R_t)$ ,

$$(S_t, I_t, R_t) \rightarrow \begin{cases} (S_t - 1, I_t + 1, R_t) & \text{w.p. } \frac{\lambda S_t}{\lambda S_t + \gamma(n-1)} \\ (S_t, I_t - 1, R_t + 1) & \text{w.p. } \frac{\gamma(n-1)}{\lambda S_t + \gamma(n-1)} \end{cases}$$

## Our Approach

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# Key Idea

- We do **not** try to **directly estimate** the value of  $I_t$  for all  $t$ .
- Instead, propose new **stopping times** ( $T_i$ ) and focus on  $I_{T_i}$ .

## Stopping Times $T_i$ and $T_{\max}$

- $T_{\max} = \min\{t \geq 0 : I_t = 0\}$ .
- $T_0 = 0$ .
- $T_i = \min\{t \geq 0 : S_t = S_0 - i\}$ .
- $\tau_i = \min\{T_i, T_{\max}\}$ .
- Since  $T_{\max} < \infty$ , we have  $\tau_i < \infty$  a.s.

## Usefulness of $\tau_i$

- Suppose  $I_{\tau_i} > 0$ . Then,  $\tau_i = T_i$  and  $S_{\tau_i} = S_0 - i$ .
- Between  $\tau_i$  and  $\tau_{i+1}$ , one and only one of the following could occur:
  1. there are a bunch of recoveries followed by an infection.
  2. all  $I_{\tau_i}$  infected people recover.
- At each recovery, the value of  $S_t$  does not change from  $S_0 - i$ .
- Recoveries between  $\tau_i$  and  $\tau_{i+1}$  is a truncated geometric random variable with parameter  $1/q_i$ , where  $q_i = 1 + \frac{1}{\mathcal{R}_0} \left( \frac{n-1}{n-I_0-i} \right)$ .
- Truncation is needed since recoveries cannot exceed  $I_{\tau_i}$ .

## Our Results

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# Expected Number of Active Infections

## Theorem

For any  $t \in \{0, \dots, S_0\}$ ,

$$\mathbb{E}[I_{\tau_t} - I_0] = \sum_{i=0}^{t-1} \left[ 1 - \frac{1}{\mathcal{R}_0} \left( \frac{n-1}{n-I_0-i} \right) \right] \mathbb{P}[T_{i+1} \leq T_{\max}].$$

# Early Termination Bounds ( $l_0 = 1$ )

## Theorem

• Suppose  $0 \leq i \leq (n - 3)/2$  and  $\zeta = 1/(\mathcal{R}_0 + 1)$ .

•  $\mathbb{P}\{T_i \geq T_{\max}\} \geq \zeta$ .

• If  $\gamma/(\gamma + \lambda) \leq 0.0654$ , then

$$\mathbb{P}\{T_i \geq T_{\max}\} \leq 2.9\zeta.$$

• If  $\gamma/(\gamma + \lambda) \leq 0.101$ , then

$$\limsup_{n \rightarrow \infty} \mathbb{P}\{T_i \geq T_{\max}\} \leq 1.38\zeta.$$

## Remarks

- If mean recovery period is 14 days, then  $\gamma = 1/14 \approx 0.071$ .
- If the mean time for one to meet others is 0.5, then  $\lambda = 2$ .
- $\zeta \approx 0.034$ .
- The bound does not depend on  $i$ .

## Early Termination Bounds ( $l_0 \geq 1$ )

### Theorem

Suppose  $\mathcal{R}_0 \geq 4$  and  $0 \leq i + l_0 \leq (n - 1)/2$ . Then,

$$\mathbb{P}\{T_i \geq T_{\max}\} \leq c_1 e^{-c_2 l_0},$$

where  $c_1 = e^{1/30}/(1 - e^{-1/30})$  and  $c_2 = 0.2/3$ .

- The bound  $c_1 e^{-c_2 l_0} \leq 1$  for  $l_0 \geq 52$ .
- The bound does not depend on  $i$ .

# Bounds for Active Infections ( $l_0 \geq 1$ )

## Theorem

• Suppose  $\mathcal{R}_0 \geq 4$  and  $0 \leq i + l_0 \leq (n - 1)/2$ .

• For  $q_j = 1 + \frac{1}{\mathcal{R}_0} \left( \frac{n-1}{n-l_0-j} \right)$ , let  $\mu_i = \sum_{j=0}^i q_j = O(i)$ .

• Let  $\epsilon \in (0, 1)$  be such that

$$\mathcal{E}_i := \left( 2[i + 1] - [1 + \epsilon]\mu_i, 2[i + 1] - [1 - \epsilon]\mu_i \right) \subseteq [2 - l_0, \infty).$$

• Then, for  $c_1, c_2$  as before and  $c_3 = \frac{2}{3} \left( 1 + \frac{\gamma}{\lambda} \ln(2) \right)$ ,

$$\mathbb{P}\{(l_{\tau_{i+1}} - l_{\tau_0}) \notin \mathcal{E}_i\} \leq c_1 e^{-c_2 l_0} + e^{-c_3 i \epsilon \left( 1 - \frac{1}{\sqrt{1+\epsilon}} \right)} + e^{-c_3 i \epsilon \left( \frac{1}{\sqrt{1-\epsilon}} - 1 \right)}.$$

- $\mathcal{E}_i \subseteq [2 - l_0, \infty)$  implies that  $l_{\tau_{i+1}} > 0$ .
- $l_{\tau_{i+1}} > 0$  implies  $l_{\tau_j} > 0$  for all  $j \in \{0, \dots, i\}$ .

## Proof Ideas

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## Early Termination Bounds

- Use [coupling](#) to show that, for  $k \geq 0$  and  $\mathcal{E}_0, \dots, \mathcal{E}_k \subset [2 - l_0, \infty)$ ,

$$\mathbb{P} \left[ \bigcap_{i=0}^k \{ (l_{\tau_{i+1}} - l_{\tau_0}) \in \mathcal{E}_i \} \right] = \mathbb{P} \left[ \bigcap_{i=0}^k \left\{ \left( (i+1) - \sum_{j=0}^i H_j \right) \in C_i \right\} \right],$$

where  $H_j \sim \text{Geom}(1/q_j)$  are independent random variables.

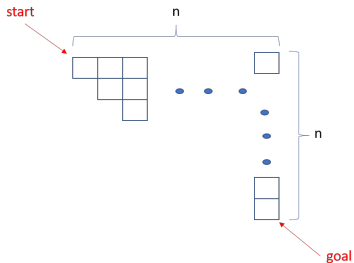
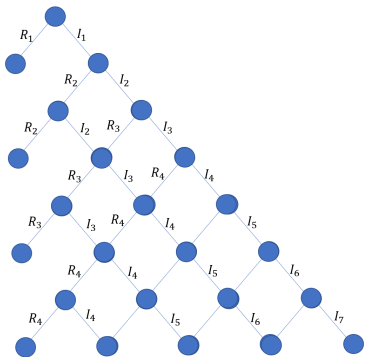
- $\mathbb{P}\{T_{k+1} < T_{\max}\} = \mathbb{P} \left[ \bigcap_{j=0}^k \{l_{\tau_{j+1}} \geq 2\} \right] \geq 1 - \sum_{i=0}^k \mathbb{P} \left\{ \sum_{j=0}^i H_j \geq i + l_0 \right\}$
- Use [bounds for sums of geometric variables](#) from [Janson '14].

## Bounds for Active Infections

For  $\mathcal{E}_i \subseteq [2 - I_0, \infty)$ ,

$$\mathbb{P}\{(I_{\tau_{i+1}} - I_{\tau_0}) \in \mathcal{E}_i\} = \mathbb{P}\left[\left\{T_i < T_{\max}\right\} \cap \left\{\left(i + 1 - \sum_{j=0}^i H_j\right) \in \mathcal{E}_i\right\}\right].$$

# Early Termination Bounds ( $l_0 = 1$ )



## Early Termination Bounds ( $l_0 = 1$ )

- Use direct counting to obtain a bound for  $\mathbb{P}\{T_{\max} = k\}$ .
- $l_{t+1} - l_t = 2(S_t - S_{t+1}) - 1$ .
- $l_t - l_0 = 2(S_0 - S_t) - t$ .
- $\{T_{\max} < T_i\} = \{S_{T_{\max}} > S_0 - i\} = \{T_{\max} < 2i + l_0\}$ .

## Expected Number of Active Infections

- Let  $X_t \equiv (S_t, I_t, R_t)$ .
- Suppose  $I_{\tau_i} > 0$ . Then, for any  $T_i \leq t < T_{i+1}$ ,

$$\mathbb{E}[I_{t+1} - I_t | X_t] = \left(1 - \frac{1}{\mathcal{R}_0} \left[ \frac{n-1}{n-I_0-i} \right]\right) \mathbb{E}[S_t - S_{t+1} | X_t].$$

- $\mathbb{E}[I_{\tau_{i+1}} - I_{\tau_i}] = \left(1 - \frac{1}{\mathcal{R}_0} \left[ \frac{n-1}{n-I_0-i} \right]\right) \mathbb{E}[S_{\tau_i} - S_{\tau_{i+1}}]$ .
- $S_{\tau_i} - S_{\tau_{i+1}} = \mathbb{I}[T_{i+1} \leq T_{\max}]$ .

- Extend the first early termination result to cover the case  $I_0 = k$ .
- Obtain bounds for  $\mathbb{E}[I_{\tau_i} - I_{\tau_0}]$  using estimates for  $\mathbb{P}\{T_i \leq T_{\max}\}$ .
- Obtain estimates for [expected time for disease extinction](#).

