

2. Let $\Omega = \{-1, +1\}^{\mathbb{N}}$ equipped with the probability denoted by \mathbb{P} , such that

$$\mathbb{P}(\{\omega \in \Omega : \pi_N(\omega) = \tilde{\omega}\}) = \frac{1}{2^N} = \mathbb{P}_N(\{\tilde{\omega}\}),$$

where $N \in \mathbb{N}$, $\tilde{\omega} \in \Omega_N = \{-1, 1\}^N$, equipped with the uniform distribution, denoted by \mathbb{P}_N and $\pi_N : \Omega \rightarrow \Omega_N$ be the canonical projection.

Consider for $k \geq 1$, $X_k : \Omega \rightarrow \{-1, 1\}$ be given by $X_k(\omega) = \omega_k$ and for $1 \leq n$, let $S_n : \Omega \rightarrow \mathbb{Z}$ be given by $S_n(\omega) = \sum_{k=1}^n X_k(\omega)$ and $S_0 = 0$.

Definition: Let \mathcal{A}_n be the events that are observable by time n . We shall say a sequence of random variables, $\{H_n\}$ is a martingale w.r.t. the filtration \mathcal{A}_n if

$$\mathbb{E}[|H_n|] < \infty \quad \text{and} \quad \mathbb{E}[H_n | S_{n-1}, \dots, S_1] = H_{n-1}$$

(a) Show that $\xi_n = S_n^2 - n$ is a martingale w.r.t the filtration \mathcal{A}_n .

(d) Let $\{V_k\}_{k \geq 1}$ be a predictable process, that is for $c \in \mathbb{R}$,

$$\{V_k = c\} \in \mathcal{A}_{k-1}.$$

Then show that

$$Z_0 = 0 \quad Z_n = \sum_{k=1}^n V_k(S_k - S_{k-1})$$

is a martingale w.r.t the filtration \mathcal{A}_n .

Solution 2(a) : We must prove that ξ_n is a martingale with respect to \mathcal{A}_n . For this , we note that

$$\mathbb{E}[|\xi_n|] \leq \mathbb{E}[|S_n|^2] + n \leq n^2 + n < \infty$$

for all n where the second inequality follows from the fact that $|S_n|^2$ is bounded above by n^2 . For the second part, fix $s_{n-1}, s_{n-2}, \dots, s_1$ such that $\mathbb{P}[S_{n-1} = s_{n-1}, S_{n-2} = s_{n-2}, \dots, S_1 = s_1] \neq 0$. Then the conditional probability and expectation with respect to this event is well defined and we have :

$$\begin{aligned} \mathbb{E}[\xi_n | S_{n-1} = s_{n-1}, \dots, S_1 = s_1] &= \sum_{k \in \mathbb{Z}} k \mathbb{P}[\xi_n = k | S_{n-1} = s_{n-1}, \dots, S_1 = s_1] \\ &= \sum_{k \in \mathbb{Z}} k \mathbb{P}[S_n^2 = n + k | S_{n-1} = s_{n-1}, \dots, S_1 = s_1] \\ &= \sum_{k \in \mathbb{Z}} k \mathbb{P}[(s_{n-1} + X_n)^2 = n + k | S_{n-1} = s_{n-1}, \dots, S_1 = s_1] \\ &= \sum_{k \in \mathbb{Z}} k \frac{\mathbb{P}[(s_{n-1} + X_n)^2 = n + k, S_{n-1} = s_{n-1}, \dots, S_1 = s_1]}{\mathbb{P}[S_{n-1} = s_{n-1}, \dots, S_1 = s_1]} \end{aligned}$$

Since X_n is independent of S_1, \dots, S_{n-1} , we get :

$$\begin{aligned}
\sum_{k \in \mathbb{Z}} k \frac{\mathbb{P}[(s_{n-1} + X_n)^2 = n + k, S_{n-1} = s_{n-1}, \dots, S_1 = s_1]}{\mathbb{P}[S_{n-1} = s_{n-1}, \dots, S_1 = s_1]} &= \sum_{k \in \mathbb{Z}} k \frac{\mathbb{P}[(s_{n-1} + X_n)^2 = n + k] \mathbb{P}[S_{n-1} = s_{n-1}, \dots, S_1 = s_1]}{\mathbb{P}[S_{n-1} = s_{n-1}, \dots, S_1 = s_1]} \\
&= \sum_{k \in \mathbb{Z}} k \mathbb{P}[(s_{n-1} + X_n)^2 = n + k] \\
&= ((s_{n-1} + 1)^2 - n) \mathbb{P}[X_n = 1] + ((s_{n-1} - 1)^2 - n) \mathbb{P}[X_n = -1] \\
&= \frac{(s_{n-1} - 1)^2 - n + (s_{n-1} + 1)^2 - n}{2} = s_{n-1}^2 - (n - 1) = \xi_{n-1}
\end{aligned}$$

as desired. Thus, ξ_n is a martingale w.r.t \mathcal{A}_n . □

Solution 2(d) : First note that $S_k - S_{k-1} = X_k$ therefore we have $Z_n = \sum_{k=1}^n V_k X_k$. Therefore :

$$\mathbb{E}[|Z_n|] \leq \sum_{k=1}^n \mathbb{E}[|V_k| |X_k|] \leq \sum_{k=1}^n \mathbb{E}[|V_k|] < \infty$$

As the V_k are bounded random variables, being functions on a finite domain. For the second condition, fix n . We know that $\{V_k = c\} \in \mathcal{A}_{k-1} \subset \mathcal{A}_{n-1}$ for all $k \leq n$. In particular, this implies that

$$\{V_k = c\} = \{(X_1, \dots, X_{n-1}) \in A'\}$$

for some set $A' \subset \{\pm 1\}^{n-1}$. But then, $X_j = S_j - S_{j-1}$, therefore $(X_1, \dots, X_{n-1}) = g(S_1, \dots, S_{n-1})$ for some function g , and it follows that

$$\{V_k = c\} = \{(S_1, \dots, S_{n-1}) \in g^{-1}(A')\}$$

for any k . In particular, for any set $B \subset D_1 \times \dots \times D_k$ where D_i denote the set of values which V_i takes for $1 \leq i \leq n$, we have

$$\{(V_1, \dots, V_k) \in B\} = \{S_1, \dots, S_{n-1} \in S'\}$$

for some $S' \subset \mathbb{Z}^n$. This establishes the following : there exists a function f' with appropriate domain and codomain such that $V_n = f'(S_1, \dots, S_{n-1})$. Using a similar argument for other $k \leq n$ this also implies that $Z_{n-1} = \sum_{k=1}^{n-1} V_k X_k = g(S_1, \dots, S_n)$ for some function g .

Let $s_{n-1}, s_{n-2}, \dots, s_1$ be such that $\mathbb{P}[S_{n-1} = s_{n-1}, \dots, S_1 = s_1] > 0$. Then we have, from definitions and from the independence of X_n from S_1, \dots, S_{n-1} , in an argument analogous to 2(a) :

$$\begin{aligned}
\mathbb{E}[Z_n | S_{n-1} = s_{n-1}, \dots, S_1 = s_1] &= \sum_K K \mathbb{P}[Z_n = K | S_{n-1} = s_{n-1}, \dots, S_1 = s_1] \\
&= \sum_K K \mathbb{P}[Z_{n-1} + V_n X_n = K | S_{n-1} = s_{n-1}, \dots, S_1 = s_1] \\
&= \sum_K K \mathbb{P}[g(s_{n-1}, \dots, s_1) + X_n f'(s_{n-1}, \dots, s_1) = K | S_{n-1} = s_{n-1}, \dots, S_1 = s_1] \\
&= \sum_K K \frac{\mathbb{P}[g(s_{n-1}, \dots, s_1) + X_n f'(s_{n-1}, \dots, s_1) = K, S_{n-1} = s_{n-1}, \dots, S_1 = s_1]}{\mathbb{P}[S_{n-1} = s_{n-1}, \dots, S_1 = s_1]} \\
&= \sum_K K \frac{\mathbb{P}[g(s_{n-1}, \dots, s_1) + X_n f'(s_{n-1}, \dots, s_1) = K] \mathbb{P}[S_{n-1} = s_{n-1}, \dots, S_1 = s_1]}{\mathbb{P}[S_{n-1} = s_{n-1}, \dots, S_1 = s_1]} \\
&= \sum_K K \mathbb{P}[g(s_{n-1}, \dots, s_1) + X_n f'(s_{n-1}, \dots, s_1) = K] \\
&= [g(s_{n-1}, \dots, s_1) + f'(s_{n-1}, \dots, s_1)] \mathbb{P}[X_n = 1] + [g(s_{n-1}, \dots, s_1) - f'(s_{n-1}, \dots, s_1)] \mathbb{P}[X_n = -1] \\
&= \frac{[g(s_{n-1}, \dots, s_1) + f'(s_{n-1}, \dots, s_1)] + [g(s_{n-1}, \dots, s_1) - f'(s_{n-1}, \dots, s_1)]}{2} \\
&= g(s_{n-1}, \dots, s_1) = Z_{n-1}
\end{aligned}$$

which shows that Z_n is a martingale. \square 3. Let X, Y, Z be discrete random variables on a probability space $(\Omega, \mathcal{F}, \mathbb{P})$. Suppose that $\mathbb{P}(Y = y) > 0$ and $\mathbb{P}(Z = z) > 0$.

(a) Show that $\mathbb{E}[XY | Y = y] = y\mathbb{E}[X | Y = y]$

Solution 3(a) : As $\mathbb{P}[Y = y] > 0$, the expression $E[XY|Y = y]$ is well defined. By definition, we have :

$$E[XY|Y = y] = \sum_{k \in \mathbb{Z}} k\mathbb{P}[XY = k|Y = y]$$

we now distinguish $y = 0$ and $y \neq 0$. Indeed, if $y = 0$ then we have :

$$\sum_{k \in \mathbb{Z}} k\mathbb{P}[XY = k|Y = 0] = \sum_{k \in \mathbb{Z}} k\mathbb{P}[k = 0|Y = 0] = 0 = yE[X|Y = y]$$

trivially. For $y \neq 0$, we have :

$$\sum_{k \in \mathbb{Z}} k\mathbb{P}[XY = k|Y = y] = \sum_{k \in \mathbb{Z}} k\mathbb{P}\left[X = \frac{k}{y} \middle| Y = y\right] = y \sum_{k \in \mathbb{Z}} \frac{k}{y} \mathbb{P}\left[X = \frac{k}{y} \middle| Y = y\right]$$

We claim that

$$\sum_{k \in \mathbb{Z}} \frac{k}{y} \mathbb{P}\left[X = \frac{k}{y} \middle| Y = y\right] = \sum_{l \in \mathbb{Z}} l\mathbb{P}[X = l|Y = y]$$

To prove this, note that if $y \nmid k$ then $\mathbb{P}\left[X = \frac{k}{y}\right] = 0$ and hence $\mathbb{P}\left[X = \frac{k}{y} \middle| Y = y\right] = 0$. Therefore, the only terms that remain on the LHS are those for which $y|k$. Reindexing with $l = \frac{y}{k}$ gives us the RHS. With this, we get :

$$y \sum_{k \in \mathbb{Z}} \frac{k}{y} \mathbb{P}\left[X = \frac{k}{y} \middle| Y = y\right] = y \sum_{l \in \mathbb{Z}} l\mathbb{P}[X = l|Y = y] = yE[X|Y = y]$$

as desired. \square