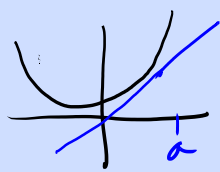


Recall :-

Kolmogorov's inequality

$h: \mathbb{R} \rightarrow \mathbb{R}$ st $\forall a \in \mathbb{R} \exists c \in \mathbb{R}$ such that $h(x) \geq h(a) + c(x-a) \quad \forall x \in \mathbb{R}$ } Convex



Jensen's inequality : $E h(x) \geq h(E(x))$ when h is convex

$\{Z_n\}_{n \geq 1}$ was a martingale then $E[h(Z_n) | Z_{n-1}, \dots, Z_1] \geq h(Z_{n-1})$

$\{Z_n\}_{n \geq 1}$ - sub-martingale - $E(Z_n) \leq Z_{n-1} \quad \forall n \geq 1$
 $E[Z_n | Z_{n-1}, \dots, Z_1] \geq Z_{n-1} \quad \forall n \geq 2$

Theorem 3-8 [Kolmogorov sub-martingale Inequality]

$\{Z_n\}_{n \geq 1}$ be a non-negative sub-martingale. Then for any positive integer n and any $a \geq 0$

$$\mathbb{P}(\max_{1 \leq i \leq n} Z_i \geq a) \leq \frac{E[Z_n]}{a} \quad \text{--- (1)}$$

Observations:

- Compare (1) with Markov inequality
- any martingale is also a sub-martingale.

Proof:- let $m \in \mathbb{N}$ be given

$J = \min \{k \geq 1 \mid Z_k \geq a\} \wedge m$
 is a bounded stopping time.

- Earlier HW
 we have seen that
 J is a stopping time.

Note: $Z_n < a \quad \forall n < m \implies J = m.$

So $Z_J \geq a$ iff $Z_n \geq a$ for some $n \leq m.$

$$\implies P(\max_{1 \leq n \leq m} Z_n \geq a) = P(Z_J \geq a)$$

$$(Z_i \geq 0) \dots \text{(Markov inequality)} \leq \frac{E[Z_J]}{a}$$

$$(Z_J = Z_{J \wedge m} \text{ as } J \text{ is bounded by } m) = \frac{E[Z_m^J]}{a} - \textcircled{+}$$

Now, $Z_m^J = Z_{J \wedge m} = \sum_{k=1}^m Z_k \mathbb{1}_{J \geq k} + Z_m \mathbb{1}_{J > m}$

$$\implies E[Z_m^J] = \sum_{k=1}^m E[Z_k \mathbb{1}_{J \geq k}] + E[Z_m \mathbb{1}_{J > m}]$$

Z_m -sub-martingale
 $(E[Z_m | \mathcal{A}_k] \geq Z_k)$

$$\leq \sum_{k=1}^m E[E[Z_m | \mathcal{A}_k] \mathbb{1}_{J \geq k}] + E[Z_m \mathbb{1}_{J > m}]$$

Conditional and (predictable) expectation

$$\text{i.e. } E[XY | \mathcal{A}_Y] = Y E[X | \mathcal{A}_Y]$$

$$= \sum_{k=1}^m E[E[Z_m \mathbb{1}_{J \geq k} | \mathcal{A}_k]] + E[Z_m \mathbb{1}_{J > m}]$$

$$E[E(X|A_x)] = E[X]$$

TP

$$\sum_{k=1}^m E[Z_m \mathbb{1}_{J=k}] + E[Z_m \mathbb{1}_{J>m}]$$

$$= E[Z_m (\sum_{k=1}^m \mathbb{1}_{J=k} + \mathbb{1}_{J>m})]$$

$$= E[Z_m] \quad - \quad (*)$$

Using $(*)$ and (\dagger) we have

$$P(\max_{1 \leq i \leq n} Z_i \geq a) \leq \frac{E[Z_n]}{a} \quad \forall n \geq 1 \quad \square$$

Recall: Corollary 3.9: let $\{Z_m\}_{m \geq 1}$ be a martingale

with $E[Z_m^2] < \infty \quad \forall m \geq 1$ then

$$P(\max_{1 \leq i \leq n} |Z_i| \geq b) \leq \frac{E[Z_m^2]}{b^2} \quad \forall m \geq 2 \quad \forall b > 0.$$

Proof: Applies Theorem 3.8 to Z_i^2 and $a = b^2$ \square

Corollary 3.10: - let $\{Z_n: n \geq 1\}$ be a non-negative martingale

$$P(\sup_{n \geq 1} Z_n \geq a) \leq \frac{E[Z_1]}{a} \quad \forall a > 0$$

Proof: $\{Z_n\}_{n \geq 1}$ is a martingale (non-negative)

Theorem 3.8 $\Rightarrow \mathbb{P}(\max_{1 \leq i \leq n} Z_i \geq a) \leq \frac{E[Z_n]}{a}$

(Earlier Corollary & Lemma 3.2) $E[Z_n] = E[Z_1] \quad \forall n \geq 1$ $= \frac{E[Z_1]}{a}$ (x)

$A_n = \{ \max_{1 \leq i \leq n} Z_i \geq a \} \Rightarrow A_n \subseteq A_{n+1} \quad \forall n \geq 1$ (xx)

$A_n \uparrow A = \bigcup_{i \geq 1} \{Z_i \geq a\}$ [key point]

Let $\epsilon_0 > 0$ be given and $0 < \epsilon_0 < a$.

$\mathbb{P}(\sup_{n \geq 1} Z_n \geq a) = \mathbb{P}(\bigcap_{\epsilon > 0} \bigcup_{i=1}^{\infty} \{Z_i \geq a - \epsilon\})$

$\leq \mathbb{P}(\bigcup_{i=1}^{\infty} \{Z_i \geq a - \epsilon_0\})$ (xxx)

Take (I): $a - \epsilon_0$ in (x) to get

$\mathbb{P}(\max_{1 \leq i \leq n} Z_i \geq a - \epsilon_0) \leq \frac{E[Z_1]}{a - \epsilon_0}$

(II):

(Take limit as $n \rightarrow \infty$ in I)

and use $A_n \uparrow A$
 $\Rightarrow \mathbb{P}(A_n) \uparrow \mathbb{P}(A)$

$\mathbb{P}(\bigcup_{i \geq 1} Z_i \geq a - \epsilon_0) \leq \frac{E[Z_1]}{a - \epsilon_0}$

(III) Use xxx to get-

$$\mathbb{P}\left(\sup_{n \geq 1} Z_n \geq a\right) \leq \frac{E[Z_1]}{a - \epsilon_0}$$

(IV) $\epsilon_0 > 0$ was arbitrary in $0 < \epsilon_0 < a$

$$\Rightarrow \mathbb{P}\left(\sup_{n \geq 1} Z_n \geq a\right) \leq \frac{E[Z_1]}{a} \quad \square$$

Theorem 3.11 (Martingale Convergence Theorem) :-

Let $\{Z_n\}_{n \geq 1}$ be a martingale and assume that $\exists M > 0$ such that $E[Z_n^2] \leq M \quad \forall n \geq 1$

Then there exists a r.v. Z such that $Z_n \rightarrow Z$ as $n \rightarrow \infty$ with probability 1.

$$\text{(i.e. } \mathbb{P}\left(\lim_{n \rightarrow \infty} Z_n = Z\right) = 1 \text{ or } \mathbb{P}(Z_n \rightarrow Z \text{ as } n \rightarrow \infty) = 1)$$

Proof:-

$\{Z_n^2\}_{n \geq 1}$ is a sub-martingale (non-negative)

$$\Rightarrow E[Z_n^2 | Z_1^2, \dots, Z_{n-1}^2] \geq Z_{n-1}^2 \quad \forall n \geq 2$$

$$\text{(TP)} \quad E[Z_n^2 | Z_1^2, \dots, Z_i^2] \geq Z_i^2 \quad \forall n > i \geq 1$$

$$\text{(Above, At } i=1 \text{ } E[Z_n^2 | Z_1^2] \geq Z_1^2)$$

Taken's expectation in \otimes

$$E[Z_n^2] \geq E[Z_{n-1}^2] \quad \forall n \geq 1$$

In particular
($E[Z_n^2]$ is increasing sequence)

$$E[Z_n^2] \geq E[Z_{n-1}^2] \quad \forall n \geq 2. \quad \text{--- } \oplus$$

\oplus and assumption that $E[Z_n^2] \leq M \quad \forall n \geq 1$
 $\exists m > 0$
(Bounded above)

$$\Rightarrow E[Z_n^2] \rightarrow \alpha \text{ (say)} \quad \text{as } n \rightarrow \infty. \quad \text{--- } \textcircled{I}$$

Next, let $k \geq 1$ and $Y_n = Z_{k+n} - Z_k \quad \forall n \geq 1$

Claim II :- $\{Y_n\}_{n \geq 1}$ is a mean 0, martingale.

Proof :- In HW 7, Book-keeping $E_x D$

$$E[Y_n^2] = E[(Z_{k+n} - Z_k)^2]$$

$$= E[Z_{k+n}^2] + E[Z_k^2] - 2E[Z_{k+n}Z_k]$$

(TP) \leftarrow

$$= E[Z_{k+n}^2] + E[Z_k^2] - 2E[E[Z_{k+n}Z_k | \mathcal{A}_k]]$$

Conditional Expectation \leftarrow
Predictable \leftarrow

$$= E[Z_{k+n}^2] + E[Z_k^2] - 2E[Z_k E[Z_{k+n} | \mathcal{A}_k]]$$

$\{Z_n\}_{n \geq 1}$ is martingale \leftarrow

$$= E[Z_{k+n}^2] + E[Z_k^2] - 2E[Z_k^2]$$

$$= E[Z_{k+n}^2] - E[Z_k^2] < \infty \quad \text{--- } \textcircled{II}$$

Apply Corollary 3.9 $\Rightarrow \forall b > 0 \in m \geq 1$

$$\mathbb{P} \left(\max_{1 \leq i \leq m} |Y_i| \geq b \right) \leq \frac{E[Y_m^2]}{b^2}$$

$$\Rightarrow \mathbb{P} \left(\max_{1 \leq i \leq m} |Z_{i+k} - Z_i| \geq b \right) \leq \frac{E[Z_{i+k}^2] - E[Z_i^2]}{b^2}$$

Let $m \rightarrow \infty$ on both sides to get

$$\forall b > 0 \quad \mathbb{P} \left(\bigcup_{i \geq 1} |Z_{i+k} - Z_i| \geq b \right) \stackrel{\textcircled{I}}{\leq} \frac{\alpha - E[Z_k^2]}{b^2}$$

- A similar argument as before in Proof of Corollary 3.10 will provide

$$\forall b > 0, 0 \leq \mathbb{P} \left(\sup_{i \geq 1} |Z_{i+k} - Z_i| \geq b \right) \leq \frac{\alpha - E[Z_k^2]}{b^2}$$

Take $k \rightarrow \infty$ on all terms

$$\forall b > 0 \quad \lim_{k \rightarrow \infty} \mathbb{P} \left(\sup_{i \geq 1} |Z_{i+k} - Z_i| \geq b \right) = 0$$

\textcircled{II} above implies $\{Z_k\}_{k \geq 1}$ is a Cauchy sequence with probability 1. [To show] [Using Borel-Cantelli]

$$\Rightarrow \exists Z \text{ st } \mathbb{P}(Z_n \rightarrow Z \text{ as } n \rightarrow \infty) = 1 \quad \square$$

Q2: HW 7 - Book-keeping Exercise : $\{Z_n\}_{n \geq 1}$ to be a martingale
 & $Y_n = Z_{k+n} - Z_k$ for some $k \geq 1$, $\forall n \geq 1$

Claim: $\{Y_n\}_{n \geq 1}$ is a mean 0 martingale

Proof:

$$\forall n \geq 1 \quad \left\{ \begin{array}{l} E|Y_n| \leq E|Z_{k+n}| + E|Z_k| < \infty \\ E[Y_n] = E[Z_{k+n}] - E[Z_k] = E[Z_1] - E[Z_1] = 0 \end{array} \right. \quad \begin{array}{l} \text{martingale} \\ \end{array}$$

let A_n - be the observable events by time n
 w.r.t. $\{Z_m\}_{m \geq 1}$

B_n - be the observable event by time n
 w.r.t. $\{Y_m\}_{m \geq 1}$

$(C \text{ is observable by time } n+k \text{ w.r.t. } Z)$

$\uparrow \uparrow \uparrow$
 $1_C = g(Z_k, \dots, Z_{n+k})$
 for some g

$(C \text{ - observable by time } n \text{ w.r.t. } Y)$

\Downarrow
 $1_C = f(Y_1, \dots, Y_n)$
 for some f

$$S_0 \quad A_{k+n} \quad \supseteq \quad B_n$$

To show: $E(Y_n | B_{n-1}) = Y_{n-1}$, it is

enough to show $E(Y_n | A_{k+n-1}) = Y_{n-1}$

$$\Leftrightarrow E(Z_{k+n} - Z_k | A_{k+n-1}) = Y_{n-1}$$

$$\Leftrightarrow E(Z_{k+n} | A_{k+n-1}) - E(Z_k | A_{k+n-1}) = Y_{n-1}$$

$\{Z_k\}$ - martingale

conditional
expectations
- predictable

$(=)$

$$Z_{k+n-1}$$

-

$$Z_k = Y_{n-1}$$

Which is true.

D

Question 1 HW 7 - Book-keeping Exercise:

$T: \Omega \rightarrow \mathbb{N}$ bc a stopping time w.r.t. martingale $\langle M_n \rangle_{n \geq 1}$

- discrete rv's -
- \mathbb{Z} -valued -

$$E[M_1 \mathbb{1}_{T > N}] \longrightarrow 0 \quad \text{as } N \rightarrow \infty.$$

Q1: $E|M_1| < \infty$ T is a r.v. such that

(M_1 - Discrete random variable)
values in \mathbb{Z}

$$\mathbb{P}(T < \infty) = 1 \quad \Rightarrow \quad E[M_1 \mathbb{1}_{T > N}] \longrightarrow 0 \quad \text{as } N \rightarrow \infty$$

Proof: $\lim_{N \rightarrow \infty} E[M_1 \mathbb{1}_{T > N}] = \sum_{k=1}^{\infty} \underbrace{\lim_{N \rightarrow \infty} \mathbb{P}(|M_1| = k, T > N)}_0 k$

let $\varepsilon > 0$ be given.

Given :- $E|M_1| = \sum_{k=1}^{\infty} k \mathbb{P}(|M_1| = k) < \infty$

Ex: $\exists N_0 : \sum_{k=N_0+1}^{\infty} k \mathbb{P}(|M_1| = k) < \varepsilon.$

$$\begin{aligned} \therefore E[M_1 \mathbb{1}_{T > N}] & \stackrel{(Ex)}{=} \sum_{k=1}^{N_0} k \mathbb{P}(|M_1| = k, T > N) + \sum_{k=N_0+1}^{\infty} k \mathbb{P}(|M_1| = k, T > N) \\ & \leq \left(\sum_{k=1}^{N_0} k \right) \mathbb{P}(T > N) + \underbrace{\sum_{k=N_0+1}^{\infty} k \mathbb{P}(|M_1| = k)}_{\leq \varepsilon} \mathbb{P}(T > N) \end{aligned}$$

$$\Rightarrow 0 < E[M_1 \mathbb{1}_{T > N}] \leq \frac{N_0(N_0+1)}{2} \mathbb{P}(T > N) + \varepsilon$$

Ex: $\exists N_1$ st $\mathbb{P}(T > N) < \frac{\varepsilon}{2(N_0)(N_0+1)} \quad \forall N > N_1$

$\exists N, \delta t$

$\forall N \geq N_1$

$$0 < E(|M_n| \mathbb{1}_{\{\tau > n\}}) < 2\varepsilon.$$

As $|E[M_n \mathbb{1}_{\tau > n}]| \leq E[|M_n| \mathbb{1}_{\tau > n}]$

$\Rightarrow \exists N_1, \delta t \quad \forall N \geq N_1$

$$|E[M_n \mathbb{1}_{\tau > n}]| < 2\varepsilon$$

$$\Rightarrow \lim_{n \rightarrow \infty} E[M_n \mathbb{1}_{\tau > n}] \rightarrow 0 \quad \square$$