

Hypothesis Tests and Threshold Crossing Probabilities in Random Walks.

- The MAP Test and the Sequential Test
- Main objective: Bounds on probability of error in each test
- Tools from Random Walks one of which is the Chernoff Bound.

i) We have a discrete signal H taking values $\{0, 1\}$, transmitted over a communication channel.

We receive noisy observations of H .

Based on data we receive, we want to decide what H is.

ii) Random Variable $Y \sim P(\cdot|H)$.

We want to test for what value H actually is. $H \in \{0, 1\}$.

Based on observations of Y , we calculate a certain statistic

Maximum A Priori Probability Test

Let H be a discrete random variable with support $\{0, 1\}$. We assign

$$p_0 = P(H=0), \quad p_1 = P(H=1).$$

$$\Rightarrow p_0 + p_1 = 1.$$

→ We want to perform a test on H where $H_0: H=0$ vs. $H_1: H=1$.

To perform this test, we make observations Y_1, Y_2, \dots, Y_n of H such that Y_1, Y_2, \dots, Y_n are iid conditional on $H=0$ are iid conditional $H=1$.

Define $\vec{y} = (y_1, y_2, \dots, y_n)$ and let $f(\cdot | \cdot)$ be the conditional density of Y given H .

$$\Rightarrow f(\vec{y} | h) = \prod_{i=1}^n f(y_i | h) \quad \text{where}$$

$$h \in \{0, 1\}.$$

From Bayes' Theorem, we know that

$$P(H=0 | \vec{y}) = \frac{p_0 f(\vec{y} | 0)}{\sum_{i=0}^1 p_i f(\vec{y} | i)}$$

$$P(H=1 | \vec{y}) = \frac{p_1 f(\vec{y} | 1)}{\sum_{i=0}^1 p_i f(\vec{y} | i)}$$

$$\frac{P(H=0 | \vec{y})}{P(H=1 | \vec{y})} = \frac{p_0 f(\vec{y} | 0)}{p_1 f(\vec{y} | 1)}$$

$$= \frac{P_0}{P_1} \prod_{i=1}^n \frac{f(y_i | 0)}{f(y_i | 1)} = \frac{P_0}{P_1} \lambda(\vec{y})$$

↓
Likelihood ratio.

The MAP test states that

if $\lambda(\vec{y}) > \frac{P_1}{P_0}$, choose $H=0$

if $\lambda(\vec{y}) \leq \frac{P_1}{P_0}$, choose $H=1$

We want to find bounds on the probability of error in the MAP test.

Why is it important to find a bound on probability of error?

→ Gives us an estimate of how useful our test is.

Define $Z_1 = \ln \frac{f(y, 0)}{f(y, 1)}$ and we

let $\gamma_1(r) = \ln \underbrace{g_{Z_1, H=1}(r)}_{\text{mgf of } Z_1 \text{ conditional on } H=1.}$

Theorem 1: Let $a = \frac{1}{n} \ln \frac{P_1}{P_0}$ and

let $\exists r_1$ such that $\gamma_1'(r_1) = a$.

Then

$$\text{i) } P(\text{choosing } H=0 \mid H=1) = P(\text{error} \mid H=1) \\ \leq e^{-n [\gamma_1(r_1) - r_1 a]}$$

$$\text{ii) } P(\text{choosing } H=1 \mid H=0) = P(\text{error} \mid H=0) \\ \leq e^{-n [\gamma_1(r_1) + (1-r_1)a]}.$$

Outline of Proof for Theorem 1:

Chernoff Bound: for any random walk
 $S_n = X_1 + X_2 + \dots + X_n$ where $\{X_i\}_{i \geq 1}$ are
iid, let $\gamma_x(r) = \ln g_x(r)$. If $\exists s$
such that $\gamma_x'(s) = a$, then
 $P(S_n \geq na) \leq e^{-n [\gamma_x(s) - sa]}.$

This is a sharp bound.

$$\text{i) } P(\text{choosing } H=0 \mid H=1) \\ = P(\lambda(\vec{y}) > \frac{P_1}{P_0} \mid H=1) \quad \text{--- } (*)$$

$$\text{We let } S_n = \ln(\lambda(\vec{y}))$$

$$\Rightarrow S_n = \sum_{i=1}^n \ln \frac{f(y_i | 1)}{f(y_i | 0)} = \sum z_i$$

S_n is a random walk because

Y_1, Y_2, \dots, Y_n are iid.

This is stated as Proposition 1 in the notes.

$$\textcircled{*} = P(S_n > na \mid H=1) \text{ where } a = \frac{1}{n} \ln \frac{p_1}{p_0}$$

From the Chernoff Bound, we have directly that

$$P(S_n > na \mid H=1) \leq e^{-n[\gamma_1(r_1) - r_1 a]}$$

This proves (i).

For (ii)

$$P(\text{choosing } H=1 \mid H=0)$$

$$= P(\chi(\vec{y})) \leq \frac{p_1}{p_0} \mid H=0$$

$$= P(S_n \leq \ln \frac{p_1}{p_0} \mid H=0)$$

$$= P(-S_n \geq -\ln \frac{p_1}{p_0} \mid H=0) \text{ --- } \textcircled{**}$$

We define $T_n \equiv -S_n$.

$$\Rightarrow T_n = \sum_{i=1}^n X_i \quad \text{s.t. } X_i = -z_i$$

(**)

$$= P(T_n \geq -na \mid H=0)$$

We use the Chernoff bound to see

$$\text{that } P(T_n \geq -na \mid H=0)$$

$$\leq e^{n[\gamma_0(r_0) - r_0(-a)]}$$

$$= e^{n[\gamma_0(r_0) + r_0 a]} \quad (***)$$

where $\gamma_0 = \ln g_{X|H=0}(r)$

and r_0 is such that $\gamma_0'(r_0) = -a$.

Claim (proved in the notes)

i) $\gamma_0(r) = \gamma_1(1-r)$ for r that lies in domain of $\gamma_2(r)$.

$$\text{ii) } r_0 + r_1 = 1$$

$$\text{iii) } \gamma_0(r_0) = \gamma_1(r_1).$$

We substitute (i), (ii) and (iii)

in (***) to get

$$P(T_n \geq -na \mid H=0)$$

$$\leq e^{n[\gamma_1(r_1) + (1-r_1)a]}.$$

This proves (ii)

□.

Chernoff Bound (Proposition 2)

For any random walk $S_n = X_1 + X_2 + \dots + X_n$
where X_i are iid, let $\gamma_x(r) = \ln g_x(r)$

If there exists s such that $\gamma_x'(s) = a$
then $P(S_n \geq na) \leq e^{n[\gamma_x(s) - sa]}$.

and this is a sharp bound.

Outline of the proof of Chernoff Bound

$$P(S_n \geq na) = P(e^{rS_n} \geq e^{rna})$$

for $r > 0$.

$$P(e^{rS_n} \geq e^{rna})$$

$$\leq \frac{E[e^{rS_n}]}{e^{rna}} \quad [\text{Markov Inequality}]$$

$$= [g_x(r)]^n e^{-rna}$$

$$\Rightarrow P(S_n \geq na) \leq [g_x(r)]^n e^{-rna}$$
$$= e^{n[\gamma_x(r) - ra]}$$

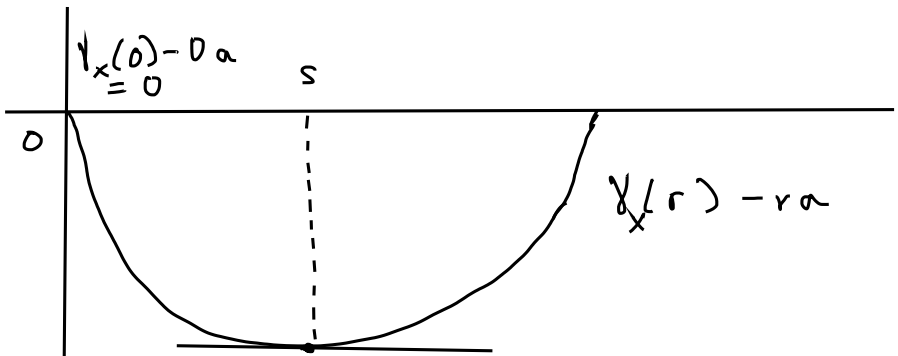
for $r > 0$.

$$\min_{r>0} \gamma_x(r) - ra$$

In theorem 3 of the notes, we use Hölder's Inequality to prove that $\gamma_x(r)$ is a convex function.

If $\gamma_x(r)$ is convex, then a local minimum is also a global minimum.

$$\gamma_x'(s) = a.$$



$$\gamma_x'(s) - a = 0.$$

The smallest value of $\gamma_x(r) - ra$ is at $r = s$.

$$\Rightarrow P(S_n \geq na) \leq e^{-n[\gamma_x(s) - sa]}$$

is the going to be the smallest bound on $P(S_n \geq na)$. \square

Sequential test

→ No fixed n .

→ $\alpha > 0$, $\beta < 0$.

J is a stopping time such that J is the smallest n for which $S_n \geq \alpha$ or $S_n \leq \beta$.

For $n \in \mathbb{N}$, the test states that if

i) If $S_n \in (\beta, \alpha)$, then record observation Y_{n+1}

ii) If $S_n \geq \alpha$ i.e. $S_J \geq \alpha$, then we stop testing and we select $H=0$

iii) If $S_n \leq \beta$ i.e. $S_J \leq \beta$, then stop testing and we select $H=1$.

Theorem 2:

$$\text{i) } P(\text{choosing } H=0 \mid H=1) = P(\text{error} \mid H=1) \\ \leq e^{-\alpha}$$

$$\text{ii) } P(\text{choosing } H=1 \mid H=0) = P(\text{error} \mid H=0) \\ \leq e^{-\beta}$$

Outline of Proof of Theorem 2

$$\text{i) } P(\text{error} \mid H=1) = P(S_J \geq \alpha \mid H=1) \\ \leq e^{-r^* \alpha} \quad \text{where}$$

$$\gamma_1(r^*) = 0.$$

Recall that $\gamma_1(r) = \ln g_{Z \mid H=1}(r)$.

In the notes I have shown that

$$\gamma_1(1) = 0. \Rightarrow r^* = 1$$

$$\Rightarrow P(S_J \geq \alpha \mid H=1) \leq e^{-\alpha}.$$

This proves (i).

$$\text{ii) } P(\text{error} \mid H=0) = P(S_J \leq -\beta \mid H=1) \\ \leq e^{-r^* \beta} = e^{-\beta}$$

□.

In section 4, I have discussed the power of the MAP test and why the sequential test is advantageous compared to MAP test.

