1. Consider two independent random variables X, Y on $(\Omega, \mathcal{F}, \mathbb{P})$.

(a) Show that the distribution of X + Y is given by the convolution of the distributions of X and Y. That is if $\mathbb{Q}(\cdot) = \mathbb{P} \circ (X + Y)^{-1}(\cdot)$, $\mathbb{P}_X(\cdot) = \mathbb{P} \circ X^{-1}(\cdot)$ and $\mathbb{P}_Y(\cdot) = \mathbb{P} \circ Y^{-1}(\cdot)$ then

$$\mathbb{Q}(A) = \int_{\mathbb{R}} \mathbb{P}_X(A - y) d\mathbb{P}_Y(dy),$$

for all Borel sets A, with $A - y = \{z \in \mathbb{R} : z + y \in A\}.$

- (b) Compute the distribution of X + Y when
 - i. $X \sim \text{Poisson}(\lambda)$ and $Y \sim \text{Poisson}(\mu)$
 - ii. $X \sim \text{Normal}(a, u)$ and $Y \sim \text{Normal}(b, v)$
 - iii. $X \sim \text{Gamma}(\alpha, \beta)$ and $Y \sim \text{Gamma}(\eta, \beta)$

Solution 1(b)(iii):We have to compute the joint distribution of X + Y when $X \sim Gamma(\alpha, \beta)$ and $Y \sim Gamma(\eta, \beta)$.

The pdf of $Gamma(\alpha, \beta)$ is

$$\mathbb{P}_X(x) = \frac{\beta^{\alpha} e^{-\beta x} x^{\alpha - 1}}{\Gamma(\alpha)}, \quad x > 0, \alpha, \beta > 0$$

and

$$\mathbb{P}_Y(y) = \frac{\beta^{\eta} e^{-\beta y} y^{\eta-1}}{\Gamma(\eta)}, \quad y > 0, \beta, \eta > 0.$$

Let Z = X + Y.

$$\begin{aligned} \mathbb{P}_{Z}(z) &= \mathbb{P}_{(X+Y)}(z) \\ &= \int_{0}^{z} \mathbb{P}(X+Y=z) \mathbb{P}(dz) \\ &= \int_{0}^{z} \mathbb{P}_{X}(x) \mathbb{P}_{y}(z-x) dx \\ &= \int_{0}^{z} \frac{x^{\alpha-1}e^{-\beta x}\beta^{\alpha}}{\Gamma(\alpha)} \cdot \frac{(z-x)^{\eta-1}e^{-\beta(z-x)}\beta^{\eta}}{\Gamma(\eta)} dx \\ &= \frac{\beta^{\alpha+\eta}e^{-\beta z}}{\Gamma(\alpha)\Gamma(\eta)} \int_{0}^{z} x^{\alpha-1}(z-x)^{\eta-1} dx \end{aligned}$$

To evaluate the integral $\int_0^z x^{\alpha-1}(z-x)^{\eta-1}dx$, we will use Gamma function integral:

$$\begin{split} \int_{0}^{z} x^{\alpha - 1} (z - x)^{\eta - 1} dx &= \int_{0}^{1} t^{\alpha - 1} z^{\alpha - 1} z^{\eta - 1} (1 - t)^{\eta - 1} z dt \quad [x = tz] \\ &= z^{\alpha + \eta - 1} \int_{0}^{1} t^{\alpha - 1} (1 - t)^{\eta - 1} dt \\ &= z^{\alpha + \eta - 1} \Gamma(\alpha, \eta) \\ &= z^{\alpha + \eta - 1} \frac{\Gamma(\alpha) \Gamma(\eta)}{\Gamma(\alpha + \eta)}. \end{split}$$

Therefore,

$$X + Y \sim Gamma(\alpha + \eta, \beta).$$

- 3. Suppose $\{X_n\}_{n\geq 1}$, X are random variables on $(\Omega, \mathcal{F}, \mathbb{P})$. Suppose $X_n \xrightarrow{d} X$ or X_n converges in distribution to X (i.e. $\mathbb{P} \circ X_n^{-1}$ converges weakly to $\mathbb{P} \circ X^{-1}$).
 - (a) Ex 5.2.11 in [?]
 - (b) Ex 5.2.19 in [?]
 - (c) Ex 5.2.23 in [?]

Solution 3(a): $X_n \xrightarrow{d} X$ then $\mathbb{P}(X_n \leq t) \to \mathbb{P}(X \leq t)$ as $n \to \infty \ \forall t \in \mathbb{R}$. Now, $\mathbb{P}(X_n^2 \leq t) = 0$ if t < 0 and $\mathbb{P}(X^2 \leq t) = 0$ if t < 0. Thus, if $t < 0 \ \mathbb{P}(X_n^2 \leq t) = 0 \to 0 = \mathbb{P}(X^2 \leq t)$. If, $t \geq 0$ and $t, \sqrt{t}, -\sqrt{t}$ are points of continuity,

$$\mathbb{P}(X_n^2 \le t) = \mathbb{X}_{\ltimes} \le \sqrt{\approx} - \mathbb{P}(X_n \le -\sqrt{t}) + \mathbb{P}(X_n = -\sqrt{t})$$
$$\to \mathbb{X} \le \sqrt{\approx} - \mathbb{P}(X \le -\sqrt{t}) + \mathbb{P}(X = -\sqrt{t}) \text{ as } n \to \infty$$
$$= \mathbb{P}(X^2 \le t).$$

Solution 3(b):Since a is a continuity point of F_X , if $\epsilon > 0$ we can select $\delta > 0$ with

$$|F_X(a-\delta) - F_X(a)| < \frac{\epsilon}{3}$$

such that $a - \delta$ is a continuity point of F_X .

Now we choose N so that $n \ge N$, this implies,

$$|F_{X_n}(a) - F_X(a)| < \frac{\epsilon}{3}$$
 and $|F_{X_n}(a-\delta) - F_X(a-\delta)| < \frac{\epsilon}{3}$.

Thus,

$$\mathbb{P}(X_n = a) \leq |F_{X_n}(a) - F_{X_n}(a - \delta)| \\
\leq |F_{X_n}(a) - F_X(a)| + |F_X(a) - F_X(a - \delta)| + |F_X(a - \delta) - F_{X_n}(a - \delta)| \\
< \epsilon$$

Thus,

$$\mathbb{P}(X_n = a) \to 0 \text{ as } n \to \infty.$$

6. Ex 10.3.4 in [?].

Solution:

Suppose $\{\mu_n\}_{n\geq 1} \Rightarrow \mu$ and $\{\mu_n\}_{n\geq 1} \Rightarrow \nu$. Then \forall bounded continuous functions $f : \mathbb{R} \to \mathbb{R}$,

$$\int f d\mu_n \to \int f d\mu \quad \text{as } n \to \infty$$

and

$$\int f d\mu_n \to \int f d\nu \quad \text{as } n \to \infty.$$

Hence $\forall \ f: \mathbb{R} \rightarrow \mathbb{R}$ with f bounded and continous,

$$\int f d\mu = \int f d\nu$$

Choose $y \in \mathbb{R}$. Now take a sequence of functions $f_n : \mathbb{R} \to \mathbb{R}$ such that $f_n \equiv 1$ on $(-\infty, y]$ and $f_n \equiv 0$ on $(y + \frac{1}{n}, \infty)$.

Then

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$$\lim_{n \to \infty} f_n = \chi_{(-\infty, y]}$$

By dominating convergence theorem,

$$\mu((-\infty, y]) = \int_{\mathbb{R}} \chi_{(-\infty, y]} d\mu$$
$$= \int \lim_{n \to \infty} f_n d\mu$$
$$= \lim_{n \to \infty} \int f_n d\mu$$
$$= \int \lim_{n \to \infty} \int f_n d\nu$$
$$= \int \lim_{n \to \infty} f_n d\nu$$
$$= \int_{\mathbb{R}} \chi_{(-\infty, y]} d\nu$$
$$= nu((-\infty, y])$$

Therefore $\mu \equiv \nu$ as $\mu(\{y\}) = \nu(\{y\}) = 0$.