

Recall ∴

$$\mathbb{P}: \mathcal{F} \rightarrow [0,1] \quad \begin{cases} \mathbb{P}(S) = 1 \\ A_1, A_2, \dots : A_i \cap A_j = \emptyset \quad i \neq j \\ \mathbb{P}(\bigcup_{i=1}^{\infty} A_i) = \sum_{j=1}^{\infty} \mathbb{P}(A_j) \end{cases}$$

- $(S, \mathcal{F}, \mathbb{P}) \equiv$ Probability space
 - ↳ set of all possible outcomes
 - ↳ Events - all subsets of S

S - countable

Probability mass function

$$X: S \rightarrow T$$

$$f_X(t) = \mathbb{P}(X=t)$$

- $X \sim \text{Binomial}(n, p)$ $X = \#$ of success in n "independent" trials

$$\mathbb{P}(X=k) = \binom{n}{k} p^k (1-p)^{n-k} \quad k=0,1,2,\dots,n$$

Expected value

Q: On an average how many success will there be in n independent Bernoulli(p) trials?

- Bernoulli(p) trials
- n -trials; Each trial has 2 outcomes
 - success (p)
 - failure ($1-p$)
 - Outcome of one trial does not affect the outcome of any other trial.

Q: $X \sim \text{Binomial}(n, p)$
What is the average value of X ?

Example :- - Roll a Die

- Possible outcomes $\equiv \{1, 2, 3, 4, 5, 6\}$
- What is the average value that shows up?

$$\frac{1 + 2 + 3 + 4 + 5 + 6}{6} = 3.5$$

Rewrite

weighted
average

$$1\left(\frac{1}{6}\right) + 2\left(\frac{1}{6}\right) + 3\left(\frac{1}{6}\right) + 4\left(\frac{1}{6}\right) + 5\left(\frac{1}{6}\right) + 6\left(\frac{1}{6}\right)$$

Probabilities with which each outcome happens

$x_1, \dots, x_n \quad \dots$

$$\frac{\sum_{i=1}^n w_i x_i}{\sum_{i=1}^n w_i}$$

$w_i = \frac{1}{6}$

$$\sum_{i=1}^6 w_i = 1$$

Expectation

Let $X : S \rightarrow T$ be a discrete random variable (so T is countable). Then the expected value (or average) of X is written as $E[X]$ and is given by

$$E[X] = \sum_{t \in T} t \cdot P(X = t)$$

provided that the sum converges absolutely. In this case we say that X has “finite expectation”.

- Refer to PSWEUR
chapter 4 for
examples

- If the sum diverges to $\pm\infty$ we say the random variable has infinite expectation.
- If the sum diverges, but not to infinity, we say the expected value is undefined.

Example 2 $\therefore X \sim \text{Binomial}(n, p)$

$$P(X=k) = \binom{n}{k} p^k (1-p)^{n-k}$$

$k=0, 1, 2, \dots, n$

$$E[X] = \sum_{k=0}^n k P(X=k)$$

$$= \sum_{k=0}^n k \binom{n}{k} p^k (1-p)^{n-k}$$

[Induction] $E_x = np$

Example : - Toss a coin till the 1st Head appears
- p - Probability of Heads

let $X \equiv$ denote the trial at which the 1st head appears

$$E[X] = ?$$

Distribution of X : $\text{Range}(X) = \mathbb{N}$

$$P(X=k) = P(\underbrace{T \dots T}_{k-1} T H) = (1-p)^{k-1} p$$

Expected value of X

$$E[X] = \sum_{k=1}^{\infty} k P(X=k) = \lim_{n \rightarrow \infty} T_n$$

where $T_n = \sum_{k=1}^n k P(X=k) =$

$$= \sum_{k=1}^n k (1-p)^{k-1} p$$

$$= p \sum_{k=1}^n k (1-p)^{k-1}$$

Ex: (Geometric series)

$\downarrow n \rightarrow \infty$

$\frac{1}{p}$

$$E[X] = \frac{1}{p}$$

Interpret $E[X]$:

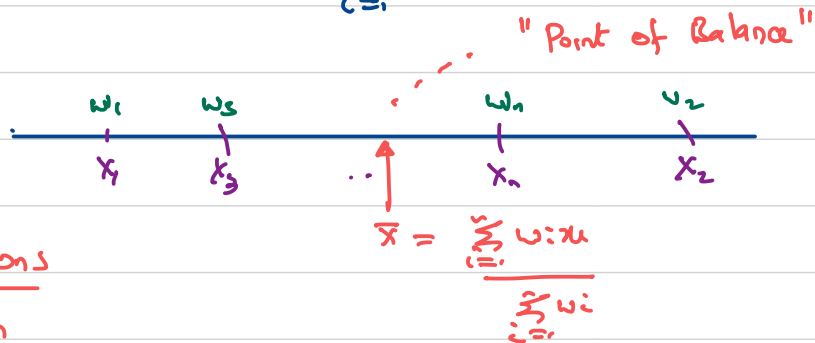
• $X \sim \text{Binomial}(10, \frac{1}{2})$

Perform trial many times, on an average we will get 5

$$E[X] = 5$$

• X_1, \dots, X_n - $\frac{\sum_{i=1}^n X_i}{n} = 3.5$

• X_1, \dots, X_n - $\frac{\sum_{i=1}^n w_i X_i}{\sum_{i=1}^n w_i} = 2$



Other notions

median

mode

Properties of Expected Value

Suppose that X and Y are discrete random variables, both with finite expected value and both defined on the same sample space S . If a and b are real numbers then

$$E[aX] = aE[X];$$

$$E[X + Y] = E[X] + E[Y]; \quad \text{and}$$

$$E[aX + bY] = aE[X] + bE[Y].$$

$$\text{If } X \geq 0 \text{ then } E[X] \geq 0.$$

} From
Definition
of
Expected value.

Variance

Let $X : S \rightarrow T$ be a discrete random variable with finite expected value. Then the variance of the random variable is written as $Var[X]$ and is defined as

$$Var[X] = E[(X - E[X])^2] = \sum_{t \in T} (t - E[X])^2 P(X = t)$$

- The standard deviation of X is written as $SD[X]$ and is defined as

$$SD[X] = \sqrt{Var[X]}$$

$$x_1, \dots, x_n \equiv \text{numbers} \quad \bar{x} = \frac{1}{n} \sum_{i=1}^n x_i$$

$$\frac{1}{n} \sum_{i=1}^n (x_i - \bar{x})^2 \equiv \text{Var}(x)$$

$$\text{Statistic: } \frac{1}{n-1} \sum_{i=1}^n (x_i - \bar{x})^2 \equiv \text{Sample variance}$$

$$\text{Var}[x] = \sum_{t \in T} \underbrace{(t - E[x])^2}_{\text{distance from mean}} \underbrace{P(X=t)}_{\text{probability}}$$

- If X has high probability of assuming t which is far from $E[X]$ then variance will be large.
- interpretation also available for small variance.

$$\text{SD}[x] = \sqrt{\text{Var}[x]}$$

units $\left\{ \begin{array}{l} X - \text{meters} \quad \dots \rightarrow \text{Var}[x] = (\text{meters})^2 \\ \text{SD}[x] = \text{meters} \end{array} \right.$

Loosely speaking : $\mu = E[x] \quad \sigma = \text{SD}[x]$

Effective Range of $x = "(\mu - \sigma, \mu + \sigma)"$

Q: $k \geq 1; \mu = E[x]; \sigma = \text{SD}[x]; \mathbb{P}(|x - \mu| \geq k\sigma) = ?$

Interpretation of Variance

- If X has a high probability of being far away from $E[X]$ the variance will tend to be large, while if X is very near $E[X]$ with high probability the variance will tend to be small.
- If we were to associate units with the random variable X (say *meters*), then the units of $Var[X]$ would be *meters*² and the units of $SD[X]$ would be *meters*.
- Informally we will view the standard deviation as a typical distance from average.
- It is possible that $Var[X]$ and $SD[X]$ could be infinite even if $E[X]$ is finite – meaning that the random variable has a clear average, but is so spread out that any finite number underestimates the typical distance of the random variable from its average.

Standardising Random Variables

- A standardized random variable X is one for which

$$E[X] = 0 \quad \text{and} \quad \text{Var}[X] = 1.$$

- Let X be a discrete random variable with finite expected value and finite, non-zero variance. Then $Z = \frac{X - E[X]}{SD[X]}$ is a standardized random variable.

$$\therefore E[Z] = E\left[\frac{X - E[X]}{SD[X]}\right] = \frac{E[X] - E[X]}{SD[X]} = 0$$

$$\text{Var}[Z] = E[Z - E[Z]]^2 = \frac{E[(X - E[X])^2]}{SD[X]^2} = 1$$

Example: $X \sim \text{Uniform } \{1, 2, 3, 4, 5, 6\}$

$$P(X=k) = 1/6 \quad k=1, 2, 3, 4, 5, 6$$

$$E[X] = \frac{1+2+3+4+5+6}{6} = 3.5$$

(In my
Sample in \mathcal{R})
~ 2.55

$$\text{Var}[X] = \sum_{k=1}^6 (k-3.5)^2 \cdot \frac{1}{6}$$

$$= \frac{(1-3.5)^2 + (2-3.5)^2 + (3-3.5)^2 + \dots}{6}$$

$$= \frac{(2.5)^2 + (1.5)^2 + (0.5)^2 + (0.5)^2 + (1.5)^2 + (2.5)^2}{6}$$

$$= 2.91666 \dots$$

(In my sample in \mathcal{R})
~ 2.82

Sampling from a given distribution

- we can use the `sample` function.
- takes a sample of the specified size (specified by `size`) from the elements of `x` using either with or without replacement (specified by `replace`). The optional `prob` argument can be used to give a vector of weights for obtaining the elements of the vector being sampled.

```
> x = c(1,2,3,4,5,6)
```

```
> probx= c(1/6,1/6,1/6,1/6,1/6,1/6)
```

```
> Rolls=sample(x, size=1800, replace=T, prob=probx)
```

Uniform(1,2,3,4,5,6)

```
> mean(Rolls)
```

```
[1] 3.501111
```

```
> var(Rolls)
```

```
[1] 3.001666
```

Sums of Rolls

Suppose we wish to simulate in R the experiment that we did in class of Rolling a die and noting down its sum. We can use the `sample`, `matrix` and `apply`.

```
> x = c(1,2,3,4,5,6)
> probx= c(1/6,1/6,1/6,1/6,1/6,1/6)
> Rolls=sample(x, size=1500, replace=T, prob=probx)
> Rollm=matrix(Rolls, 5)
> # above creates a matrix 5 columns and 300 Rows
> Rollsums = apply(Rollm, 2, sum)
```

Sums of Rolls

