

Ishaan Taneja

**Grading:**

30 marks- Complete submission of Problem 1,3,5,7

70 marks- Problem 3

From Probability and Statistics with Examples Using R

## 1. Exercise 3.2.4

**Solution: 1**

(a)  $\text{Range}(X) = \{0, 1, 2\}$

$\text{Range}(Y) = \{0, 1, 2\}$

Y/X	X = 0	X = 1	X = 2	P(Y = y)
Y=0	1/12	0	3/12	4/12
Y=1	2/12	1/12	0	3/12
Y=2	3/12	1/12	1/12	5/12
P(X=x)	6/12	2/12	4/12	

Therefore, marginal distribution of X is:

$$P(X=0) = \frac{6}{12} \quad ; \quad P(X=1) = \frac{2}{12} \quad \text{and,} \quad P(X=2) = \frac{4}{12}$$

And, marginal distribution of Y is:

$$P(Y=0) = \frac{4}{12} \quad ; \quad P(Y=1) = \frac{3}{12} \quad \text{and,} \quad P(Y=2) = \frac{5}{12}$$

(b) Conditional distribution of X given Y=2 is:

$$P(X=x|Y=2) = \frac{P(X=x \cap Y=2)}{P(Y=2)} \quad ; \quad x=0,1,2$$

Therefore,

$$\begin{aligned} P(X=0|Y=2) &= \frac{P(X=0 \cap Y=2)}{P(Y=2)} \\ &= \frac{3/12}{5/12} \\ &= \frac{3}{5} \end{aligned}$$

$$\begin{aligned} P(X=1|Y=2) &= \frac{P(X=1 \cap Y=2)}{P(Y=2)} \\ &= \frac{1/12}{5/12} \end{aligned}$$

$$= \frac{1}{5}$$

$$\begin{aligned} P(X = 2|Y = 2) &= \frac{P(X = 2 \cap Y = 2)}{P(Y = 2)} \\ &= \frac{1/12}{5/12} \\ &= \frac{1}{5} \end{aligned}$$

(c) Conditional distribution of Y given X=2 is:

$$P(Y=y|X=2) = \frac{P(Y = y \cap X = 2)}{P(X = 2)} ; y=0,1,2$$

Therefore,

$$\begin{aligned} P(Y = 0|X = 2) &= \frac{P(Y = 0 \cap X = 2)}{P(X = 2)} \\ &= \frac{3/12}{4/12} \\ &= \frac{3}{4} \end{aligned}$$

$$\begin{aligned} P(Y = 1|X = 2) &= \frac{P(Y = 1 \cap X = 2)}{P(X = 2)} \\ &= \frac{0}{4/12} \\ &= 0 \end{aligned}$$

$$\begin{aligned} P(Y = 2|X = 2) &= \frac{P(Y = 2 \cap X = 2)}{P(X = 2)} \\ &= \frac{1/12}{4/12} \\ &= \frac{1}{4} \end{aligned}$$

(d) The random variables X and Y are independent if and only if

$$P(X = x \cap Y = y) = P(X = x)P(Y = y) ; x \in \text{Range}(X), y \in \text{Range}(Y)$$

$$\begin{aligned} P(X = 0, Y = 0) &= \frac{P(X = 0 \cap Y = 0)}{P(Y = 0)} \\ &= \frac{1/12}{4/12} \\ &= \frac{1}{4} \end{aligned}$$

And,

$$P(X = 0)P(Y = 0) = \left(\frac{6}{12}\right) \left(\frac{4}{12}\right)$$

$$= \frac{1}{6}$$

$$\neq P(X = 0 \cap Y = 0)$$

Hence, X and Y are not independent random variables.

2. Exercise 3.2.5

**Solution: 2**

$$\text{Range}(X) = \{0, 1\}$$

$$P(X=0) = \frac{1}{3} \text{ and } P(X=1) = \frac{2}{3}$$

$$\text{Range}(Y) = \{0, 1, 2\}$$

$$P(Y=0) = \frac{1}{5}, P(Y=1) = \frac{1}{5} \text{ and } P(Y=2) = \frac{3}{5}$$

Since, X and Y are independent random variables.

$$\therefore P(X = x \cap Y = y) = P(X = x)P(Y = y)$$

The joint distribution of X and Y is:

Y/X	X = 0	X = 1	P(Y = y)
Y=0	1/15	2/15	1/5
Y=1	1/15	2/15	1/5
Y=2	3/15	6/15	3/5
P(X=x)	1/3	2/3	

3. Exercise 3.2.9

**Solution: 3**

N= Number of earthquakes in a year

M= Number of earthquakes in a year with magnitude at least 5

$N \sim \text{Poisson}(\lambda)$

$$P(N=n) = \frac{e^{-\lambda} \lambda^n}{n!} ; n=0,1,2,\dots,\infty$$

And,

$(M|N = n) \sim \text{Binomial}(n, p)$

$$P(M=m|N=n) = {}^n C_m p^m (1-p)^{n-m} ; m=0,1,2,\dots,n$$

(a) The joint distribution of M and N is:

$$P(M = m \cap N = n) = P(M = m|N = n)P(N = n)$$

$$= {}^n C_m p^m (1-p)^{n-m} e^{-\lambda} \frac{\lambda^n}{n!}$$

$$= \frac{n!}{m!(n-m)!} p^m (1-p)^{n-m} e^{-\lambda} \frac{\lambda^n}{n!}$$

$$= \frac{1}{m!(n-m)!} p^m (1-p)^{n-m} e^{-\lambda} \lambda^n$$

$$\therefore P(M = m \cap N = n) = \begin{cases} \frac{1}{m!(n-m)!} p^m (1-p)^{n-m} e^{-\lambda} \lambda^n & ; \text{ for } m = 0,1,2,\dots,n \text{ and } n=m,m+1,\dots,\infty \\ 0 & ; \text{ otherwise} \end{cases}$$

(b) The marginal distribution of M is determined by:

$$P(M = m) = \sum_{n=m}^{\infty} P(M = m \cap N = n)$$

$$\begin{aligned}
&= \sum_{n=m}^{\infty} \frac{1}{m!(n-m)!} p^m (1-p)^{n-m} e^{-\lambda} \lambda^n \\
\therefore P(M=m) &= \frac{1}{m!} e^{-\lambda} (\lambda p)^m \sum_{n=m}^{\infty} \frac{\lambda^{n-m}}{(n-m)!} (1-p)^{n-m} ; m > 0
\end{aligned}$$

(c) Put  $k=n-m$

If  $n = m \implies k = n - m = 0$

and, if  $n = \infty \implies k = n - m = \infty$

Hence,

$$P(M=m) = \frac{1}{m!} e^{-\lambda} (\lambda p)^m \sum_{k=0}^{\infty} \frac{\lambda^k}{k!} (1-p)^k ; m > 0$$

(d) Using the infinite series equality  $e^x = \sum_{k=0}^{\infty} \frac{x^k}{k!}$ , we get:

$$\begin{aligned}
P(M=m) &= \frac{1}{m!} e^{-\lambda} (\lambda p)^m e^{\lambda(1-p)} \\
&= \frac{1}{m!} e^{-\lambda(1-(1-p))} (\lambda p)^m \\
\therefore P(M=m) &= \frac{1}{m!} e^{-(\lambda p)} (\lambda p)^m ; m > 0
\end{aligned}$$

Hence,  $M \sim \text{Poisson}(\lambda p)$

#### 4. Exercise 3.3.7

**Solution: 4**

Let  $X \sim \text{Geometric}(p)$  and  $Y \sim \text{Geometric}(p)$  be independent.  $Z=X+Y$

(a)  $\text{Range}(X)=\{1,2,3,\dots\}$   
 $\text{Range}(Y)=\{1,2,3,\dots\}$   
 $\implies \text{Range}(Z)=\{2,3,4,\dots\}$

(b)

$$\begin{aligned}
P(Z=n) &= P(X+Y=n) \\
&= P(\cup_{i=1}^{n-1} (X=i, Y=n-i)) \\
&= \sum_{i=1}^{n-1} P(X=i, Y=n-i) \quad (\text{Mutually exclusive events}) \\
&= \sum_{i=1}^{n-1} P(X=i)P(Y=n-i) \quad (X \text{ and } Y \text{ are independent random variables}) \\
&= \sum_{i=1}^{n-1} p(1-p)^{i-1} \times p(1-p)^{n-i-1} \\
&= \sum_{i=1}^{n-1} p^2(1-p)^{n-2} \\
&= p^2(1-p)^{n-2} \sum_{i=1}^{n-1} 1
\end{aligned}$$

$$\therefore P(Z = n) = (n - 1)p^2(1 - p)^{n-2} ; n = 2, 3, 4, \dots$$

(c) For n=2,

$$P(Z = 2) = (2 - 1)p^2(1 - p)^{2-2}$$

$$\therefore P(Z = 2) = p^2$$

For n=3,

$$P(Z = 3) = (3 - 1)p^2(1 - p)^{3-2}$$

$$\therefore P(Z = 3) = 2p^2(1 - p)$$

Now, we have to obtain the value of 'p' such that:

$$\begin{aligned} P(Z = 3) &> P(Z = 2) \\ 2p^2(1 - p) &> p^2 \\ p^2(2 - 2p - 1) &> 0 \\ \text{Since, } p^2 &> 0 \\ \implies 2 - 2p - 1 &> 0 \\ \therefore p &< \frac{1}{2} \end{aligned}$$

Hence, for all  $p < 0.5$ , the  $P(Z=3)$  is larger than  $P(Z=2)$ .

5. Exercise 3.3.11

**Solution: 5**

$X_1, X_2, X_3$  and  $X_4$  are i.i.d Bernoulli(p)

$$Y = X_1 + X_2 \sim \text{Binomial}(2, p)$$

$$Z = X_3 + X_4 \sim \text{Binomial}(2, p)$$

(a)

$$\begin{aligned} P(Y = y \cap Z = z) &= P(X_1 + X_2 = y \cap X_3 + X_4 = z) \\ &= \sum_{x_1 \in \text{Range}(X_1)} \sum_{x_3 \in \text{Range}(X_3)} P(X_1 = x_1 \cap X_2 = y - x_1 \cap X_3 = x_3 \cap X_4 = z - x_3) \\ &= \sum_{x_1 \in \text{Range}(X_1)} \sum_{x_3 \in \text{Range}(X_3)} P(X_1 = x_1)P(X_2 = y - x_1)P(X_3 = x_3)P(X_4 = z - x_3) \end{aligned}$$

Because,  $X_1, X_2, X_3, X_4$  are independent random variables

$$\begin{aligned} \therefore P(Y = y \cap Z = z) &= \sum_{x_1 \in \text{Range}(X_1)} P(X_1 = x_1)P(X_2 = y - x_1) \sum_{x_3 \in \text{Range}(X_3)} P(X_3 = x_3)P(X_4 = z - x_3) \\ &= P(Y = x_1 + x_2)P(Z = x_3 + x_4) \\ &= P(Y = y)P(Z = z) \end{aligned}$$

Hence, Y and Z are independent random variables. Now, the joint distribution of Y and Z is:

$$\begin{aligned} P(Y = y \cap Z = z) &= P(Y = y)P(Z = z) \\ &= {}^2C_y p^y (1 - p)^{2-y} \times {}^2C_z p^z (1 - p)^{2-z} \\ &= {}^2C_y {}^2C_z p^{y+z} (1 - p)^{4-(y+z)} \end{aligned}$$

Y/Z	Z = 0	Z = 1	Z = 2
Y=0	$(1-p)^4$	$2p(1-p)^3$	$p^2(1-p)^2$
Y=1	$2p(1-p)^3$	$4p^2(1-p)^2$	$2p^3(1-p)$
Y=2	$p^2(1-p)^2$	$2p^3(1-p)$	$p^4$

- (b) Y and Z are independent if and only if  
 $P(Y = y \cap Z = z) = P(Y = y)P(Z = z)$ ;  $y \in \text{Range}(Y)$  and  $z \in \text{Range}(Z)$   
 For  $y=0$  and  $z=0$ :

$$\begin{aligned} P(Y = 0)P(Z = 0) &= {}^2C_0 p^0 (1-p)^2 \times {}^2C_0 p^0 (1-p)^2 \\ &= (1-p)^4 \\ \therefore P(Y = 0)P(Z = 0) &= P(Y = 0 \cap Z = 0) \end{aligned}$$

Similarly, for other values of Y and Z, it can be shown that:  
 $P(Y = y \cap Z = z) = P(Y = y)P(Z = z)$ .

- (c) The theorem 3.3.6 states that:  
 Let  $n > 1$  be a positive integer. For each  $j \in \{1, 2, \dots, n\}$  define a positive integer  $m_j$  and suppose  $X_{i,j}$  is an array of mutually independent random variables for  $j \in \{1, 2, \dots, n\}$  and  $i \in \{1, 2, \dots, m_j\}$ . Let  $f_j$  be functions such that the quantity

$$Y_j = f_j(X_{1,j}, X_{2,j}, \dots, X_{m_j,j})$$

is defined for the outputs of the  $X_{i,j}$  variables. Then the resulting variables  $Y_1, Y_2, \dots, Y_n$  are mutually independent.

Here,

$$Y_1 = Y = X_1 + X_2$$

$$Y_2 = Z = X_3 + X_4$$

$\therefore$  Y and Z are independent random variables.

6. Exercise 3.3.15

**Solution: 6**

$X \sim \text{Geometric}(p)$

$Y \sim \text{Geometric}(q)$

X and Y are independent random variables. (given)

$Z = \min(X, Y)$

**METHOD I:**

- (a) Let us define the events:

$$A = \{X = n \text{ and } Y = n\}$$

$$B = \{X = n \text{ and } Y > n\}$$

$$C = \{X > n \text{ and } Y = n\}$$

Event  $\{Z = n\}$  occurs if either A or B or C occurs.

Also, the events A, B and C disjoint events as:

$$\{A \cap B\} = \phi ; \{A \cap C\} = \phi ; \{B \cap C\} = \phi \text{ and } \{A \cap B \cap C\} = \phi$$

Therefore, the event  $\{Z = n\}$  can be written as disjoint union:

$$\{Z = n\} = \{X = n, Y = n\} \cup \{X = n, Y > n\} \cup \{X > n, Y = n\}$$

- (b)  $P(Z = n) = P(X = n, Y = n) + P(X = n, Y > n) + P(X > n, Y = n)$

Because, disjoint union

Since, X and Y are independent random variables.

$$\begin{aligned} P(Z = n) &= P(X = n)P(Y = n) + P(X = n)P(Y > n) + P(X > n)P(Y = n) \\ &= (p(1-p)^{n-1}) \times (q(1-q)^{n-1}) + (p(1-p)^{n-1}) \times (1-q)^n + (1-p)^n \times (q(1-q)^{n-1}) \\ \therefore P(Z = n) &= [(1-p)(1-q)]^{n-1}(pq + p(1-q) + (1-p)q) \end{aligned}$$

(c)

$$\begin{aligned} P(Z = n) &= [(1-p)(1-q)]^{n-1}(pq + p(1-q) + (1-p)q) \\ &= [(1-p)(1-q)]^{n-1}[1 - (1-p)(1-q)] \\ &= (1-r)^{n-1}r \end{aligned}$$

Therefore,  $Z \sim \text{Geometric}(r)$  where,  $r=1-(1-p)(1-q)$

### METHOD II:

$A_j \sim \text{Bernoulli}(p)$  ;  $j=1,2,\dots$

$B_k \sim \text{Bernoulli}(q)$  ;  $k=1,2,\dots$

Also,  $A_j$  and  $B_k$  variables collectively are mutually independent.

X=Number of first  $A_j$  that produces a result of 1

Y=Number of first  $B_k$  that produces a result of 1

(a)

$$C_j = \begin{cases} 1 & ; \text{if either } A_j = 1 \text{ or } B_j = 1 \text{ or both} \\ 0 & ; A_j=0=B_j \end{cases}$$

Now,

$$\begin{aligned} P(C_j = 1) &= P(A_j = 1, B_j = 0) + P(A_j = 0, B_j = 1) + P(A_j = 1, B_j = 1) \\ &= P(A_j = 1)P(B_j = 0) + P(A_j = 0)P(B_j = 1) + P(A_j = 1)P(B_j = 1) \\ &= p(1-q) + (1-p)q + pq \\ \therefore P(C_j = 1) &= 1 - (1-p)(1-q) \end{aligned}$$

And,

$$\begin{aligned} P(C_j = 0) &= P(A_j = 0, B_j = 0) \\ &= P(A_j = 0)P(B_j = 0) \\ \therefore P(C_j = 0) &= (1-p)(1-q) \end{aligned}$$

Hence,

$$P(C_j = c) = \begin{cases} 1 - (1-p)(1-q) & ; \text{for } C_j = 1 \\ (1-p)(1-q) & ; \text{for } C_j = 0 \end{cases}$$

Therefore,  $C_j \sim \text{Bernoulli}(r)$  where,  $r=1-(1-p)(1-q)$

(b) Since,  $A_j$  and  $B_k$  variables collectively are mutually independent and  $C_j$  depends on  $A_j$  and  $B_j$  only.

Hence, the sequence  $C_1, C_2, \dots$  are mutually independent random variables.

- (c) The random variable  $Z$  represents the number of first  $C_j$  that results in a 1.  
 The random variable  $C_j = 1$  if either  $A_j = 1$  or  $B_j = 1$  or both.  
 Also,  $X$ =Number of first  $A_j$  that produces a result of 1, and  $Y$ =Number of first  $B_k$  that produces a result of 1.  
 Hence,  $\{Z = n\} = \{X = n, Y = n\} \cup \{X = n, Y > n\} \cup \{X > n, Y = n\}$   
 $\implies Z = \min\{X, Y\}$

- (d)  $P(Z = n) = P(X = n, Y = n) + P(X = n, Y > n) + P(X > n, Y = n)$   
 Because, disjoint union  
 Since,  $X$  and  $Y$  are independent random variables.

$$\begin{aligned} P(Z = n) &= P(X = n)P(Y = n) + P(X = n)P(Y > n) + P(X > n)P(Y = n) \\ &= (p(1-p)^{n-1}) \times (q(1-q)^{n-1}) + (p(1-p)^{n-1}) \times (1-q)^n + (1-p)^n \times (q(1-q)^{n-1}) \\ &= [(1-p)(1-q)]^{n-1}(pq + p(1-q) + (1-p)q) \\ &= [(1-p)(1-q)]^{n-1}(pq + p(1-q) + (1-p)q) \\ &= [(1-p)(1-q)]^{n-1}[1 - (1-p)(1-q)] \\ \therefore P(Z = n) &= (1-r)^{n-1}r \end{aligned}$$

Therefore,  $Z \sim \text{Geometric}(r)$  where,  $r=1-(1-p)(1-q)$

#### 7. Exercise 4.4.3

##### Solution: 7

Let  $X$  be the random variable representing the return on investment and let  $A, B,$  and  $C$  represent the events that the economy will be stronger, the same, and weaker in the next quarter, respectively.

We are given that:

$$E(X|A)=3 \quad ; \quad E(X|B)=1 \quad ; \quad E(X|C)=-1$$

$$P(A)=0.1 \quad ; \quad P(B)=0.4 \quad ; \quad P(C)=0.5$$

$$S.D(X|A)=3 \quad ; \quad S.D(X|B)=2 \quad ; \quad S.D(X|C)=3$$

Therefore,

$$V(X|A)=9 \quad ; \quad V(X|B)=4 \quad ; \quad V(X|C)=9$$

Now,

$$\begin{aligned} E(X) &= E(X|A)P(A) + E(X|B)P(B) + E(X|C)P(C) \\ &= (3 \times 0.1) + (1 \times 0.4) + ((-1) \times 0.5) \\ \therefore E(X) &= 0.2 \end{aligned}$$

$$\begin{aligned} V(X) &= (V(X|A) + (E(X|A))^2)P(A) + (V(X|B) + (E(X|B))^2)P(B) + (V(X|C) + \\ &\quad (E(X|C))^2)P(C) - (E(X))^2 \\ &= (9 + 3^2) \times 0.1 + (4 + 1^2) \times 0.4 + (9 + (-1)^2) \times 0.5 - (0.2)^2 \\ &= 1.8 + 2 + 5 - 0.04 \\ \therefore V(X) &= 8.76 \end{aligned}$$

#### 8. Exercise 4.4.4

##### Solution: 8



Let A be the event that a standard light bulb is selected from the box and B be the event that a Super D-Lux light bulb is selected from the box.

$$P(A)=0.9 \quad ; \quad P(B)=0.1$$

Let, X be the random variable that represent the lifetime of a bulb.

$$E(X|A)=4 \quad ; \quad S.D(X|A)=1 \implies \text{Var}(X|A)=1$$

$$E(X|B)=8 \quad ; \quad S.D(X|B)=3 \implies \text{Var}(X|B)=9$$

Now,

$$\begin{aligned} E(X) &= E(X|A)P(A) + E(X|B)P(B) \\ &= (4 \times 0.9) + (8 \times 0.1) \\ \therefore E(X) &= 4.4 \end{aligned}$$

$$\begin{aligned} V(X) &= (V(X|A) + (E(X|A))^2)P(A) + (V(X|B) + (E(X|B))^2)P(B) - (E(X))^2 \\ &= (1 + 4^2) \times 0.9 + (9 + 8^2) \times 0.1 - (4.4)^2 \\ &= 15.3 + 7.3 - 19.36 \\ \therefore V(X) &= 3.24 \end{aligned}$$

Hence,

$$\begin{aligned} S.D(X) &= \sqrt{\text{Var}(X)} \\ &= \sqrt{3.24} \\ \therefore S.D(X) &= 1.8 \end{aligned}$$