History of the stick breaking representation for the Dirichlet process

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Summary

- What is the stick breaking construction?
- Stick breaking construction almost in Blackwell and McQueen (1973)

- Stick breaking construction in Ferguson (1973) when combined with McCloskey (1965)
- Sethuraman's stick breaking construction

A nonparametric prior is just a probability distribution on \mathcal{P} the space of all probability measures (say on the real line). Measurable sets in \mathcal{P} are of the form $\{P : P(A) < r\}$. So we should specify the distribution of $(P(A_1), P(A_2), \ldots, P(A_k))$, etc. Ferguson (1973) defined the Dirichlet process $\mathcal{D}(\alpha, \beta)$ to the random probability measure for which

 $(P(A_1), P(A_2), \ldots, P(A_k)) \sim Dirich(\alpha\beta(A_1), \alpha\beta(A_2), \ldots, \alpha\beta(A_k))$

for all partitions (A_1, A_2, \ldots, A_k) of the real line.

A nonparametric prior can also be defined as the distribution of a random variable P taking values in \mathcal{P} .

The stick breaking construction does just this.

Let $\mathbf{V} = (V_1, V_2, ...)$ be i.i.d. $Beta(1, \alpha)$. Let $p_1 = V_1, p_2 = (1 - v_1)V_2, ...$ Then $\mathbf{p} = (p_1, p_2, ...)$ is a random discrete distribution. Let $\mathbf{Z} = (Z_1, z_2, ...)$ be i.i.d. β and be independent of \mathbf{V} . Let

$$P(A) = \sum_{1}^{\infty} p_i \delta_{Z_i}(A).$$

This P is a random probability measure and it defines a nonparametric prior. It is the stick breaking representation of the Dirichlet process.

The class of all nonparametric priors are the same as the class of all exchangeable sequences of random variables!

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Let X_1, X_2, \ldots be an infinite sequence of exchangeable (def?) sequence of random variables with a joint distribution Q.

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Then, from De Finetti's theorem

1. The empirical distribution functions $F_n(x) \to F(x)$ with probability 1 for all x.

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 The empirical distribution functions F_n(x) → F(x) with probability 1 for all x. In fact, sup_x |F_n(x) - F(x)| → 0 with probability 1. (Note that F(x) is a random distribution function.)

2. The empirical probability measures P_n converge to a random probability measure P weakly with probability 1.

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- The limit P of the empirical probability measures of X₁, X₂,... is also the limit of the empirical probability measures of X₂, X₃,.... Thus the distribution of P given X₁ (the posterior distribution) is the distribution of P under Q_{X1} and, by mere notation, is v^{Qx1}.

The Pólya urn sequence is an example of an infinite exchangeable random variables.

Let β be a pm on R_1 and let $\alpha > 0$. Define the joint distribution $Pol(\alpha, \beta)$ of X_1, X_2, \ldots through

$$X_1 \sim eta(\cdot), \ X_2 | X_1 \sim rac{lpha eta(\cdot) + \delta_{X_1}(\cdot)}{lpha + 1}$$

$$X_n|(X_1,\ldots,X_{n-1})\sim \frac{\alpha\beta(\cdot)+\sum_1^{n-1}\delta_{X_i}(\cdot)}{\alpha+n-1}, n=3,4,\ldots$$

This defines $Pol(\alpha, \beta)$ as an exchangeable probability measure. (It takes just some effort to establish this.)

What about the distribution of $(X_2, X_3, ...)|X_1$?

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What about the distribution of $(X_2, X_3, ...)|X_1$? It is $Pol(\alpha + 1, \frac{\alpha\beta + \delta_{X_1}}{\alpha + 1})$.

• The nonparametric prior $\nu^{Pol(\alpha,\beta)}$ is the same as the Dirichlet prior $\mathcal{D}(\alpha\beta)$!

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- That is, the distribution of (P(A₁),..., P(A_k)) for any partition (A₁,..., A_k), under Pol(α, β), is the finite dimensional Dirichlet D(αβ(A₁),..., αβ(A_k)). This is proved in Blackwell and MacQueen (1973).

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For any A, $P(A) \sim Beta(\alpha\beta(A), \alpha\beta(A^c))$. Can we allow $A = \{X_1\}$ in the above?

• The conditional distribution of $(X_2, X_3, ...)$ given X_1 is $Pol(\alpha + 1, \frac{\alpha\beta + \delta_{X_1}}{\alpha + 1})$.

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- Though each *P_n* is a discrete rpm and the limit *P* in general will be just a rpm.
- For the present case of a Pólya urn sequence, Blackwell and MacQueen (1973) show that P({X₁,...,X_n}) → 1 with probability 1 and thus P is a discrete rpm. (A little tricky. We will show some details.)

Dirichlet prior based on a Pólya urn sequences The conditional distribution of P given X_1 is $\mathcal{D}(\alpha\beta + \delta_{X_1})$.

 $B(\alpha\beta(\{X_1\})+1,\alpha\beta(R_1\setminus\{X_1\})).$

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The conditional distribution of $P(\{X_1, \ldots, X_n\})$ given (X_1, \ldots, X_n) is $Beta(\alpha\beta(\{X_1, \ldots, X_n\}) + n, \alpha\beta(R_1 \setminus \{X_1, \ldots, X_n\}))$ and $E(P(\{X_1, \ldots, X_n\} | X_1, \ldots, X_n)) = \frac{\alpha\beta(\{X_1, \ldots, X_n\}) + n}{\alpha + n}$

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From now on, assume that β is non-atomic.

The above conditional distribution of $P(\{X_1\})$ given X_1 becomes $B(1, \alpha)$ which does not depend on X_1 and thus X_1 and $P(\{X_1\})$ are independent.

Let Y_1, Y_2, \ldots be the distinct values among X_1, X_2, \ldots listed in the order of their appearance.

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Consider the sequence X_2, X_3, \ldots and remove all occurrences of X_1 which is the same as Y_1 .

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and all these are independent of $Y_1, Y_2, Y_3...$ which are i.i.d. β .

Since P is discrete and just sits on the set $\{X_1, X_2, ...\}$ which is $\{Y_1, Y_2, ...\}$,

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However, note that the statement of the stick breaking construction does not assume any properties of β !

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Further more $\frac{\mathbf{p}^{-R}}{1-\rho_R} \sim \frac{\mathbf{p}^{-1}}{1-\rho_1} \sim \mathbf{p}$, and so on.

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Ferguson showed that

$$\mathcal{D}^* = \sum_1^\infty \pi_i^* \delta_{Z_i}$$

has distribution $\mathcal{D}(\alpha\beta)$ if **Z** are i.i.d. β and $(\pi_1^*, \pi_2^*, ...)$ are independent of **Z** and are the ordered normalized jumps of a Gamma process on [0, 1] with shape parameter α .

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McCloskey (1965) showed that a size biased permutation of $(\pi_1^*, \pi_2^*, \dots)$ is $GEM(\alpha)$ and so we can rewrite

$$P=\sum_{1}^{\infty}p_{i}\delta_{Z_{i}}$$

and get the stick breaking construction.

Sethuraman (1994)



Let $\alpha > 0$ and let $\beta(\cdot)$ be a pm on \mathcal{X} .

We do not assume that β is non-atomic. Restrictions like $\mathcal{X} = R_1$ do not have to made.

Let V_1, V_2, \ldots , be i.i.d. $B(1, \alpha)$ and let Z_1, Z_2, \ldots be independent of V_1, V_2, \ldots and be i.i.d. $\beta(\cdot)$.

Let $p_1 = V_1, p_2 = (1 - V_1)V_2, p_3 = V_3(1 - V_1)(1 - V_2), \dots$

The stick breaking construction is

$$P(\cdot) = P(\mathbf{p}, \mathbf{Z})(\cdot) = \sum_{1}^{\infty} p_i \delta_{Z_i}(\cdot)$$

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The stick breaking construction is

$$P(\cdot) = P(\mathbf{p}, \mathbf{Z})(\cdot) = \sum_{1}^{\infty} p_i \delta_{Z_i}(\cdot)$$

It is clearly a discrete random probability measure. We have the special identity

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where \mathbf{p}^{-1} , \mathbf{Z}^{-1} have the obvious meanings. We could have split the above with index R, (even a random index R) instead of the index 1. We will use this identity to prove that the distribution of P is $\mathcal{D}(\alpha\beta)$ and to obtain the posterior distribution.

The special identity shows that

$$P = p_1 \delta_{Z_1} + (1 - p_1) P^*$$

where all the random variables are independent, $p_1 \sim B(1, \alpha), Z_1 \sim \beta$ and the two rpm's P, P^* have the same distribution.

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where all the random variables are independent, $p_1 \sim B(1, \alpha), Z_1 \sim \beta$ and the two rpm's P, P^* have the same distribution.

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In Sethuraman (1994) we show that $\mathcal{D}(\alpha\beta)$ is a solution to this equation, and also that, if there is a solution then it is unique.

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What about the posterior distribution?

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Let *R* be a random variable such $Q(R = r | \mathbf{p}) = p_r, r = 1, 2, ...$ and let $Y = Z_R$. Then

$$Q(Y \in A|P) = Q(Y \in A|(\mathbf{p}, \mathbf{Z}))$$

=
$$\sum_{r} Q(Y \in A, R = r|(\mathbf{p}, \mathbf{Z}))$$

=
$$\sum_{r} Q(Z_r \in A)p_r = P(A)$$

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Thus Y is a like an observation from P and we need the distribution of P given Y.

The special identity gives

$$P = \rho_R \delta_Y + (1 - \rho_R) P(\mathbf{p}^{-R}/(1 - \rho_R), \mathbf{Z}^{-R}).$$

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Conditional on (R, Y), the right hand side has distribution

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Thus the distribution of *P* given *Y* is $\mathcal{D}(\alpha\beta + \delta_Y)$.
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 $\mathcal{D}(\alpha\beta)$ is not itself a discrete probability measure.